### Learning Stochastic Differential Equations

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Fundamental problem in many engineering apps.

- Signal processing
- Quality control
- Control theory



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#### **Classical setting**

- Data = i.i.d. samples from  $P_{\theta}(x)$
- $\theta$  = low-dimensional parameter
- $n = \text{large value compared to } \dim(\theta)$



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#### **Classical setting**

E.g.: Learn  $\theta$  in  $y_{\ell} = \langle x_{\ell}, \theta \rangle + \varepsilon_{\ell}$  from i.i.d.  $\{(y_{\ell}, x_{\ell})\}_{\ell=0}^{n}$ 

Data = trajectory of a stochastic diff. eq. (SDE)

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$$\{x(t)\}_{t=0}^T : \quad dx(t) = F(x(t);\theta)dt + db(t)$$

 $\theta$  = very large graph, G = (V, E)

E.g.: *G* = interaction between components of *x* 



$$dx_{i}(t) = -mx_{i}(t)dt + \sum_{j \in \partial i} (x_{j}(t) - x_{i}(t))dt + db_{i}(t)$$

Data = trajectory of a stochastic diff. eq. (SDE)

$$\{x(t)\}_{t=0}^T : \quad dx(t) = F(x(t);\theta)dt + db(t)$$

 $\theta$  = very large graph

E.g.: *G* = interaction between components of *x* 



*T* = minimum value to recover graph with prob.  $\approx$  1

# **Graphs & SDEs** Example: Gene regulatory networks



Yeast



'18srRnaa'	'biod5'	'THR5'	'YER022w5'	'YAL058W/CNE1'	'YAL042W/'	'YAL031C/FUN21'
'18srRnab'	'biodm'	'THR3'	'YER022wM'	'YAL058C-a/'	'YAL043C-a/'	'YAL030W/
'18srRnac'	'biod3'	'TRP5'	'YER022w3'	'YAL056W/'	'YAL041W/CDC24'	SNC1_ex1'
'18srRnad'	'cre5'	'TRPM'	'YAL069W/'	'YAL055W/'	'YAL040C/CLN3'	'YAL030W/
'18srRnae'	'crem'	'TRP3'	'YAL067C/SEO1'	'YAL054C/ACS1'	'YAL039C/CYC3'	SNC1_ex2'
'25srRnaa'	'cre3'	'DAP5'	'YAL066W/'	'YAL053W/'	'YAL038W/CDC19'	'YAL029C/MYO4'
'25srRnab'	'25srRnad'	'DAPM'	'YAL065C/'	'YAL051W/'	'YAL037W/'	'YAL028W/'
'25srRnac'	'25srRnae'	'DAP3'	'YAL065C-a/'	'YAL049C/'	'YAL036C/'	'YAL027W/'
'BIOB5'	'LYSA5'	'YFL039C5'	'YAL064W/FLO9_i'	'YAL048C/'	'YAL035W/FUN12'	'YAL026C/DRS2'
'biobm'	'LYSAM'	'YFL039CM'	'YAL063C/'	'YAL047C/'	'YAL034W-a/'	'YAL025C/MAK16'
'biob3'	'LYSA3'	'YFL039C3'	'YAL062W/GDH3'	'YAL046C/'	'YAL035C-a/'	'YAL024C/LTE1'
'bioc5'	'PHE5'	'YER148w5'	'YAL061W/'	'YAL045C/'	'YAL034C/FUN19'	'YAL023C/PMT2'
'biocm'	'PHEM'	'YER148wM'	'YAL060W/'	'YAL044C/GCV3'	'YAL033W/FUN53'	'YAL022C/'
'bioc3'	'PHE3'	'YER148w3'	'YAL059W/SIM1'	'YAL043C/PTA1'	'YAL032C/FUN20'	'YALO21C/CCR4'

Yeast

#### ≈ 6000 genes



'18srRnaa'	'biod5'	'THR5'	'YER022w5'	'YAL058W/CNE1'	'YAL042W/'	'YAL031C/FUN21'
'18srRnab'	'biodm'	'THR3'	'YER022wM'	'YAL058C-a/'	'YAL043C-a/'	'YAL030W/
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'18srRnad'	'cre5'	'TRPM'	'YAL069W/'	'YAL055W/'	'YAL040C/CLN3'	'YAL030W/
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'bioc3'	'PHE3'	'YER148w3'	'YAL059W/SIM1'	'YAL043C/PTA1'	'YAL032C/FUN20'	'YAL021C/CCR4'

Yeast

# ≈ 800 genes for cell cycle [Spellman 98]



Nodes = genes Edges = interaction (activation/repression)





Yeast

#### Gene expression t. series [Stanford yeast cell cycle database ]

 $x_i(t) = i^{\text{th}}$  gene expression level at time t



Yeast

E.g.: Chen 05 et al. uses SDEs to study yeast cell-cycle



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$$A + B \stackrel{k',k}{\longleftrightarrow} C + D$$
$$dx_A = \left(k' x_C x_D - k x_A x_B\right) dt + db$$

Yeast



Simple model for  
gene interaction  
$$dx_{i} = \begin{bmatrix} \theta_{i,0} & \theta_{i,1} & \theta_{i,2} & \dots & \theta_{i,12} & \theta_{i,23} & \dots \end{bmatrix} \begin{bmatrix} 1 \\ x_{1} \\ x_{2} \\ \dots \\ x_{1}x_{2} \\ x_{2}x_{3} \\ \dots \end{bmatrix} dt + db_{i}$$

 $\theta_{i,11} \neq 0 \Leftrightarrow \mathbf{2}$  and  $\mathbf{3}$  interact to regulate *i* 

$$dx_{i} = \begin{bmatrix} \theta_{i,0} & \theta_{i,1} & \theta_{i,2} & \dots & \theta_{i,12} & \theta_{i,23} & \dots \end{bmatrix} \begin{bmatrix} 1 \\ x_{1} \\ x_{2} \\ \dots \\ x_{1}x_{2} \\ x_{2}x_{3} \\ \dots \end{bmatrix} dt + db_{i}$$

Learning  $supp(\theta)$ 

Learning network of gene interactions

$$dx_{i} = \begin{bmatrix} \theta_{i,0} & \theta_{i,1} & \theta_{i,2} & \dots & \theta_{i,12} & \theta_{i,23} & \dots \end{bmatrix} \begin{bmatrix} 1 \\ x_{1} \\ x_{2} \\ \dots \\ x_{1}x_{2} \\ x_{2}x_{3} \\ \dots \end{bmatrix} dt + db_{i}$$

#### Problem:

- 1. How to fit this (or other SDEs models) to data?
- 2. How much data do we need?



$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\}$$

$$T_{Alg,G} = \inf \{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \} \sim \mathbf{0} ??$$

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$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim p ??$$

Quantity of interest

$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim 0 ??$$

$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim p ??$$

**Fact:** Learning a sparse  $\theta$  in  $y_{\ell} = \langle x_{\ell}, \theta \rangle + \varepsilon_{\ell}$  from

i.i.d. pairs  $\{y_{\ell}, x_{\ell}\}_{\ell=0}^{n}$  is possible if  $n \sim \log p$ 

$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim 0 ??$$

$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim p ??$$

$$T_{Alg,G} = \inf \left\{ T : \mathbf{P}(\hat{G} = G) > 1 - \delta \right\} \sim \log p ??$$

### Linear SDEs & graphical models

In linear SDEs,

$$dx_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) dt + db_i(t)$$

 $\theta$  encodes the interaction among components of x

# Linear SDEs & graphical models

Learning SDEs is related to the broader problem of learning a graphical model from data



$$\mathrm{d}x_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) \mathrm{d}t + \mathrm{d}b_i(t)$$

#### Stationary distribution

 $x(t) \sim N(0, \Sigma)$  $\theta \Sigma + \Sigma \theta^{T} + I = 0$ 

$$\mathrm{d}x_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) \mathrm{d}t + \mathrm{d}b_\mathrm{i}(t)$$

Stationary distribution

 $x(t) \sim N(0, \Sigma)$  $\theta \Sigma + \Sigma \theta^{T} + I = 0$ 

Given *n* i.i.d samples from  $N(0, \Sigma)$  estimate  $\hat{G} = \operatorname{supp}(\Sigma^{-1})$ 

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### Learning Gaussian graphical models

- Friedman, Hastie, & Tibshirani 08
- Meinshausen & Buhlmann 06

# Given *n* i.i.d samples from $N(0, \Sigma)$ estimate $\hat{G} = \operatorname{supp}(\Sigma^{-1})$

### Why is our work different?

- $\Sigma$  has less information than  $\theta$
- We have  $\{x(t)\}$ , not i.i.d. samples

# Given *n* i.i.d samples from $N(0, \Sigma)$ estimate $\hat{G} = \operatorname{supp}(\Sigma^{-1})$

#### $\Sigma$ has less information than $\theta$

$$\Theta_1 = \begin{bmatrix} -2 & 1 & 0 \\ -1 & -2 & 1 \\ 0 & -1 & -2 \end{bmatrix} \qquad \Theta_2 = \begin{bmatrix} -2 & -1 & -1 \\ 1 & -2 & -1 \\ 1 & 1 & -2 \end{bmatrix} \qquad \Sigma = \begin{bmatrix} 1/4 & 0 & 0 \\ 0 & 1/4 & 0 \\ 0 & 0 & 1/4 \end{bmatrix}$$

The system of SDEs is linear in  $\boldsymbol{\theta}$ 

$$dx_{i}(t) = \sum_{j} \theta_{ij} x_{j}(t) dt + db_{i}(t)$$
$$Y = X \theta_{i} \qquad \varepsilon$$

$$Y = \begin{bmatrix} \Delta x_i^{(1)} \\ \vdots \\ \Delta x_i^{(n)} \end{bmatrix} \qquad \qquad X = \begin{bmatrix} x_1^{(1)} & \cdots & x_p^{(1)} \\ \vdots & \ddots & \vdots \\ x_1^{(n)} & \cdots & x_p^{(n)} \end{bmatrix}$$

The system of SDEs is linear in  $\boldsymbol{\theta}$ 

$$dx_{i}(t) = \sum_{j} \theta_{ij} x_{j}(t) dt + db_{i}(t)$$
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Given *Y*, *X* satisfying  $Y = X \theta_i + \varepsilon$  estimate  $\partial i = \operatorname{supp}(\theta_i)$ 

# Given *Y*, *X* satisfying $Y = X \theta_i + \varepsilon$ estimate $\partial i = \operatorname{supp}(\theta_i)$

### L<sub>1</sub> - regularized least squares

- Tibshirani 96
- Zhao & Yu 06
- Wainwright 09

# Given *Y*, *X* satisfying $Y = X \theta_i + \varepsilon$ estimate $\partial i = \operatorname{supp}(\theta_i)$

#### Why is our work different?

• Rows of *Y* and *X* obtained from {*x*(*t*)} are not i.i.d.

# **Underlying challenge:** Spaced data or correlation?



	# samples	correlation
$\eta$ decreases	1	1
$\eta$ increases	$\downarrow$	$\downarrow$

# Prior work: SDEs

Start with

$$dx_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) dt + db_i(t)$$

Compute the likelihood function



Estimate  $\theta$  from

 $\underset{\Theta \in \Re^{p \times p}}{\operatorname{argmax}} L(\Theta; \{x(t)\}_{t=0}^{T})$ 

### ML methods for SDEs

- Basawa & Rao 80
- Kutoyants 04

# **Prior work:** SDEs

Start with

$$\mathrm{d}x_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) \mathrm{d}t + \mathrm{d}b_i(t)$$

Compute approx. likelihood function  $L(\Theta; \{x(\eta \ell)\}_{\ell=1}^n)$ 

Estimate  $\theta$  from

 $\underset{\Theta \in \mathbb{R}^{p \times p}}{\operatorname{argmax}} L(\Theta; \{x(\eta \ell)\}_{\ell=1}^n)$ 

### Approximate ML methods for sampled SDEs

- Dacunha-Castelle & Florens 86
- Pederson 95
- AitSahalia 02

# Prior work: SDEs

Start with

$$dx_i(t) = \sum_{j=1}^p \theta_{ij} x_j(t) dt + db_i(t)$$

Compute the likelihood function

Estimate  $\theta$  from

$$\underset{\Theta \in \Re^{p \times p}}{\operatorname{argmax}} L(\Theta; \{x(t)\}_{t=0}^{T})$$

 $L(\Theta; \{x(t)\}_{t=0}^{T})$ 

### Why is this work different?

- Focus is on low dimensional setting
- ML is not a selection algorithm ( $\theta \neq G$ )
- Guarantees only hold asymptotically in T

# The algorithm

Consider the non-linear SDE

$$\mathrm{d}x_i(t) = \sum_{j=1}^{p'} \theta_{ij} g_j(x(t)) \mathrm{d}t + \mathrm{d}b_i(t)$$

Example



 $dx_1 = (\theta_{11}g_1(x_3) + \theta_{12}g_2(x_1, x_2))dt + db_1$ 

# The algorithm

The SDE is linear in  $\theta$  : use a regression

$$\mathrm{d}x_i(t) = \sum_{j=1}^{p'} \theta_{ij} g_j(x(t)) \mathrm{d}t + \mathrm{d}b_i(t)$$

# The algorithm

The SDE is linear in  $\theta$  : use a regression

$$dx_{i}(t) = \sum_{j=1}^{p'} \theta_{ij}g_{j}(x(t))dt + db_{i}(t)$$
Naïve derivation
$$\hat{\theta}_{i} = \underset{\Theta \in \mathbb{R}^{p}}{\operatorname{argmin}} \int_{0}^{T} \left( dx_{i}(t) - \sum_{j=1}^{p'} \Theta_{j}g_{j}(x(t))dt \right)^{2} + \lambda \|\Theta\|_{1}$$

# **The algorithm:** $RLS(\lambda)$

The SDE is linear in  $\theta$  : use a regression

$$dx_{i}(t) = \sum_{j=1}^{p'} \theta_{ij}g_{j}(x(t))dt + db_{i}(t)$$

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$$\underbrace{\operatorname{Doing the math right}}_{\Theta \in \mathbb{R}^{p}} \left( db\right)^{2} \to dt$$

$$\operatorname{argmin}_{\Theta \in \mathbb{R}^{p}} \frac{1}{2T} \int_{0}^{T} \left( \sum_{j=1}^{p'} \Theta_{j}g_{j}(x) \right)^{2} dt - \frac{1}{T} \int_{0}^{T} \sum_{j=1}^{p'} \Theta_{j}g_{j}(x) dx_{i} + \lambda \|\Theta\|_{1}$$

# **The algorithm:** $RLS(\lambda)$

$$\hat{\theta}_{i} = \underset{\Theta \in \mathbb{R}^{p}}{\operatorname{argmin}} \frac{1}{2T} \int_{0}^{T} \left( \sum_{j=1}^{p'} \Theta_{j} g_{j}(x) \right)^{2} dt - \frac{1}{T} \int_{0}^{T} \sum_{j=1}^{p'} \Theta_{j} g_{j}(x) dx_{i} + \lambda \|\Theta\|_{1}$$

$$\mathcal{L}_{i}(\Theta; \{x(t)\}_{t=0}^{T})$$

#### What is the **sample-complexity**?

$$T_{\text{RLS},G} = \inf \left\{ T : \mathbf{P}(\text{supp}(\hat{\theta}) = \text{supp}(\theta)) > 1 - \delta \right\}$$

# **Characterization of** $T_{\text{RLS}}$



Given G = (V, E) define

$$dx_i = -mx_i dt + \sum_{j \in \partial i} (x_j - x_i) dt + db_i$$

Let G be sparse of maximum degree  $\Delta$ 

Theorem If  $p, m \gg \Delta$  then  $C(\Delta)m \log p < T_{\text{RLS},G} < C'(\Delta)m^2 \log p$ 

-Bento, Ibrahimi, Montanari, Proceedings of Neural Information Processing Systems (NIPS), 2010 -Bento, Ibrahimi, Montanari, IEEE International Symposium on Information Theory (ISIT), 2011

# **Characterization of** T<sub>RLS</sub>



Given G = (V, E) define

$$dx_i = -mx_i dt + \sum_{j \in \partial i} (x_j - x_i) dt + db_i$$

Now, *G* can be dense  $(\Delta \sim p)$ 

Theorem If p, m >> 1 then  $Cmp < T_{RLS,G} < C'mp$ 

-Bento, Ibrahimi, Montanari, IEEE Transactions in information Theory, 2013 -Bento, Ibrahimi, Montanari, IEEE International Symposium on Information Theory (ISIT), 2011

# **General behaviour**

- *T* ~ log *p*: Large sparse systems can be learned from few samples
- $T \sim m$ : Fast systems are harder to learn



# **General characterization of** T<sub>RLS</sub>

General theorem for  $dx_i(t) = \sum_{j=1}^{p} \theta_{ij} x_j(t) dt + db_i(t)$ requires some assumptions

In particular, it requires assumptions similar to the ones in Zhao & Yu 06 for sparse linear regression

- Entries smallest value is lower bounded
- Strong stability assumption
- Restricted convexity assumption
- Irrepresentable condition

# **Important ideas behind proof**

 Work with discretized SDE and take limits
 Proof structure similar to Zhao & Yu 06 requires two new concentration bounds

$$\mathbb{P}\left(\left|\hat{G}_{ij}\right| > \epsilon\right) \le C_1 e^{-C_2 n\eta \epsilon^2} \qquad \mathbb{P}\left(\left|\hat{Q}_{ijk} - Q^0\right| > \epsilon\right) \le C_1' e^{-C_2' n\eta \epsilon^2} \\ \hat{G}_{ij} = \frac{\mathrm{d}\mathcal{L}(\theta_i; \{x(k\eta)\}_{k=0}^n)}{\mathrm{d}\theta_{ij}} \qquad \hat{Q}_{ijk} = \frac{\mathrm{d}^2 \mathcal{L}(\theta_i; \{x(k\eta)\}_{k=0}^n)}{\mathrm{d}\theta_{ij} \mathrm{d}\theta_{ik}}$$

unlike in Zhao & Yu 06, G and Q are not of the form  $C^{\dagger}Z$  for a random vector Z with i.i.d. entries but are of the form  $Z^{\dagger}RZ$  for some matrix R.

# Follow up work

• Autoregressive processes

$$x(t+1) = \sum_{r=0}^{k-1} \Theta(r) x(t-r) + u(t)$$

Bolstad et al. 2011 (group Lasso + tailored incoherence condition)

• Hidden variables (observe x)

$$x(t+1) = \Theta_0 x(t) + \Theta_1 h(t) + u_1(t)$$
  
$$h(t+1) = \Theta_2 x(t) + \Theta_3 h(t) + u_2(t)$$

A. Jalali & S. Sanghavi 2012 (only want  $\Theta_0$  + L1 + nuclear norm penalties + two incoherence cond. )

# Follow up work

• Non-parametric

Ruttor et al. 2013 (estimate drift + sampled paths + Gaussian Process + approximate EM)

$$\mathrm{d}x(t) = F(x(t))\mathrm{d}t + D^{1/2}\mathrm{d}b(t)$$

Jung et al. 2014 (learn dependencies in GP + Blackman-Tukey estimator + group Lasso OR graphical Lasso + thresholding); (IT lower-bound )

$$(k,l) \notin E \text{ iff } [\mathbf{S}^{-1}(\omega)]_{k,l} = 0 \ \forall \ \omega \in [0,1)$$

$$\mathbf{S}(\omega) = \sum_{i=-\infty}^{\infty} R(t) e^{-j2\pi t\omega}; R(t) = \mathbb{E}(x(t)x^{\dagger}(0))$$

### **Extension:** Non-linear SDEs

1. Algorithm analysis only applies for linear SDEs

2. Some non-linear SDEs can be described as a linear combination of a set of basis functions

### **Extension:** Non-linear SDEs

1. Algorithm analysis only applies for linear SDEs

2. Some non-linear SDEs can be described as a linear combination of a set of basis functions

#### Open problem

Does the same analysis hold for non-linear SDEs?

$$dv(t) = -\gamma v(t)dt - \nabla U(q(t))dt + \sigma db(t)$$
  

$$dq(t) = v(t)dt$$
  

$$U(q) = \frac{1}{2} \sum_{(i,j)} C_{ij}^{0} \left( \left\| q^{(i)} - q^{(j)} \right\| - D_{ij}^{0} \right)^{2}$$



$$dv(t) = -\gamma v(t)dt - \nabla U(q(t))dt + \sigma db(t)$$
  

$$dq(t) = v(t)dt$$
  

$$U(q) = \frac{1}{2} \sum_{(i,j)} C_{ij}^{0} \left( \left\| q^{(i)} - q^{(j)} \right\| - D_{ij}^{0} \right)^{2}$$



Drift = linear combination of basis functions

$$\left\{v_{i}(t)\right\}_{i \in [p]}, \left\{q_{i}(t) - q_{j}(t)\right\}_{i,j \in [p]}, \left\{\frac{q_{i}(t) - q_{j}(t)}{\left\|q_{i}(t) - q_{j}(t)\right\|}\right\}_{i,j \in [p]}$$







Sample-complexity for learning regular graphs of different sizes *p* 



 $T_{RLS} \sim \log p$ 

Sample-complexity for learning regular graphs of different sizes *p* 



# Summary

- 1. For fast (large m) linear SDEs, RLS has optimal sample complexity (log p or p)
- 2. Empirical results suggest RLS has optimal scaling of  $T_{\rm RLS}$  with p for sparse non-linear SDEs
- 3. Upper bound on  $T_{RLS}$  suggests that performance of RLS degrades for slow (small *m*) linear SDEs
- 4. For fast linear SDEs  $T_{\rm RLS} \sim$  between *m* and  $m^2$

# Thank you