Behind Markov Chain Monte-Carlo.

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Plan

Motivation

Scaling Adaptation

Multidimensional Scaling

Adaptative Metropolis-Hastings Algorithm

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An Application

Some theoretical results

Motivation

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MCMC allow to simulate any probability distribution π (typically, large dimensional space)...

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- Today, Monte-Carlo methods have become a basic tool for inference in complex stochastic models on large datasets.
- On the top of that, such analysis are often done routinely allowing only limited expert supervision Require to find methods to tune the parameters automatically !

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Metropolis-Hastings Algorithm

▶ Propose a move Y_{n+1} from a transition kernel with density $q(X_n, \cdot)$.

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- ▶ Propose a move Y_{n+1} from a transition kernel with density $q(X_n, \cdot)$.
- Accept the move with probability $\alpha(X_n, Y_{n+1})$ where

$$\alpha(x,y) = 1 \land \frac{\pi(y)q(y,x)}{\pi(x)q(x,y)}.$$

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► If the move is accepted, set X_{n+1} = Y_{n+1}; otherwise, stay at the current position X_{n+1} = X_n.

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Metropolis Algorithm

▶ $Y_{k+1} = X_k + Z_{k+1}$ where $Z_{k+1} \sim_{i.i.d.} q$, and q is symmetric, q(z) = q(-z)

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- ► In this case, q(x, y) = q(y, x) and the acceptance rate does not depend on the proposal distribution

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▶ ... biased random walk where some moves are rejected.

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Scaling



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Diffusive Limits

► Stationary distribution: $\pi^{(d)}(x_1, \ldots, x_d) = \prod_{i=1}^d f(x_i)$ on \mathbb{R}^d (asymptotic = $d \to \infty$)

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• Metropolis proposal: $q_{\theta}^{(d)}(x_1, \ldots, x_d) \sim \mathcal{N}\left(0, (\theta^2/d)\mathbf{I}_d\right)...$ with variance decreasing as 1/d.

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- Interpolated process: Z_t^(d) = X_{[td],1}^(d), we consider a single component and we speed up the time scale by d.

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- Interpolated process: Z_t^(d) = X_{[td],1}^(d), we consider a single component and we speed up the time scale by d.
- When d becomes large, a single component basically see the mean of the others (mean-field)...

Scaling Adaptation



Figure: Diffusive limits for different values of d

Scaling Adaptation

Diffusive Limits

▶ $Z^{(d)} \Rightarrow Z$, where Z solves the Langevin SDE

$$dZ_t = v^{1/2}(\theta)dB_t + (1/2)v(\theta)\nabla \log f(Z_t)dt$$
$$v(\theta) = 2\theta^2 \Phi\left(-\theta\sqrt{I}/2\right)$$

where Φ is the distribution function of $\mathcal{N}(0,1)$ and

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▶ $v(\theta)$ is the speed of the diffusion: $Z_t = \tilde{Z}_{v(\theta)t}$ where $\{\tilde{Z}_t\}$ is a solution of the normalized Langevin SDE

$$d\tilde{Z}_t = dB_t + (1/2)\nabla \log f(\tilde{Z}_t)dt.$$

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Scaling Adaptation

Speed / Acceptance rate

Mean Acceptance rate (stationary regime)

$$\tau^{(d)}(\theta) = \iint \pi^{(d)}(\mathbf{x}) q_{\theta}^{(d)}(\mathbf{y} - \mathbf{x}) \left\{ 1 \wedge \frac{\pi^{(d)}(\mathbf{y})}{\pi^{(d)}(\mathbf{x})} \right\} d\mathbf{x} d\mathbf{y} \; .$$

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Result: τ^(∞)(θ) = lim_{d→∞} τ^(d)(θ) exists and it is possible to relate the speed of the diffusion to the mean acceptance rate !

$$v(\theta) = \tau^{(\infty)}(\theta) \left\{ \Phi^{-1}(\tau^{(\infty)}(\theta)/2) \right\}^2$$

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► The speed is optimal for the value θ_* of the parameter which satisfies $\tau^{(\infty)}(\theta_*) = \bar{\tau} \approx 0.234...$

Scaling Adaptation

How to control the Acceptance Rate

• Objective: Finding the scaling factor θ solving

$$h(\theta) \stackrel{\text{def}}{=} \iint \alpha(\mathbf{x}, \mathbf{y}) q_{\theta}(\mathbf{y} - \mathbf{x}) \pi(\mathbf{x}) d\mathbf{x} d\mathbf{y} - \bar{\tau} = 0,$$

where $\alpha(\mathbf{x}, \mathbf{y}) = \{1 \wedge \pi(\mathbf{y}) / \pi(\mathbf{x})\}.$

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- ▶ Nevertheless, denoting θ_k the scaling value at iteration k, $\alpha(X_k, Y_{k+1}) - \overline{\tau}$ may be seen as a "noisy" observation of $h(\theta_k)...$

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- Nevertheless, denoting θ_k the scaling value at iteration k, α(X_k, Y_{k+1}) − τ̄ may be seen as a "noisy" observation of h(θ_k)...
- Suggest to use a stochastic approximation procedure to tune θ.

Scaling Adaptation

Controlled Metropolis Algorithm

Proposition & Accept/Reject

$$\begin{split} Y_{k+1} &= X_k + \theta_k \mathcal{N}(0, \text{Id}) \\ X_{k+1} &= \begin{cases} Y_{k+1} & \text{with prob. } \alpha(X_k, Y_{k+1}) \\ X_k & \text{otherwise} \end{cases} \end{split}$$

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Update the scaling factor

 $\theta_{k+1} = \theta_k + \gamma_{k+1} \left\{ \alpha(X_k, Y_{k+1}) - \bar{\tau} \right\}$

where $\lim_{k\to\infty} \gamma_k = 0$ and $\sum_{k=1}^{\infty} \gamma_k = \infty$.

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Scaling Adaptation



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-Multidimensional Scaling

Multidimensional scaling

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 $\pi_{\Sigma_d}^{(d)}(\mathbf{x}) = |\Sigma_d|^{-1} \pi^{(d)} \left(\Sigma_d^{-1} \mathbf{x} \right), \quad \pi^{(d)}(x_1, \dots, x_d) = \prod_{i=1}^a f(x_i)$ $q \sim N(0, (\sigma^2/d) \mathrm{Id})$

then $Z_t^{(d)} = X_{[td],1}$ converges to the solution a Langevin SDE.

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$$\lim_{d} \frac{d^{-1} \sum_{i=1}^{d} \lambda_{d,i}^2}{\left(d^{-1} \sum_{i=1}^{d} \lambda_{d,i}\right)^2}$$

where $\lambda_{d,i}$ eigenvalues of Σ_d .

-Multidimensional Scaling

Adaptive MCMC with multidim. scaling

1. Simulate

$$\begin{split} Y_{k+1} &= X_k + \mathcal{N}(0, \sigma_k \Gamma_k) \\ X_{k+1} &= \begin{cases} Y_{k+1} & \text{with proba. } \alpha(X_k, Y_{k+1}) \\ X_k & \text{otherwise} \end{cases} \end{split}$$

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2. Update the target mean and covariance

$$\mu_{k+1} = \mu_k + \gamma_{k+1} (X_{k+1} - \mu_k)$$

$$\Gamma_{k+1} = \Gamma_k + \gamma_{k+1} \left\{ (X_{k+1} - \mu_k) (X_{k+1} - \mu_k)^{\mathrm{T}} - \Gamma_k \right\}$$

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3. Control the global scale of the proposal

$$\sigma_{k+1} = \sigma_k + \gamma_{k+1} \left(\alpha(X_k, Y_{k+1}) - \bar{\tau} \right)$$

-Multidimensional Scaling



Figure: d = 12, $\pi \sim \mathcal{N}(0, \Gamma)$, $\operatorname{cond}(\Gamma) \approx 100$, $q \sim \mathcal{N}(0, (2.32^2/d) \mathrm{I})$

-Multidimensional Scaling



Figure: d = 12, $\pi \sim \mathcal{N}(0, \Gamma)$, cond(Γ) ≈ 100 , $q \sim \mathcal{N}(0, 2.32^2/d\Gamma)$

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Figure: d = 12, $\pi \sim \mathcal{N}(0, \Gamma)$, $\operatorname{cond}(\Gamma) \approx 100$, $q \sim \mathcal{N}(0, \sigma_k \Gamma_k)$, with adaptive multidimensional scaling

Multidimensional Scaling

Tricks and Improvements

No need to estimate the covariance matrix at each iteration [batch means = OK]

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- No need to estimate the covariance matrix at each iteration [batch means = OK]
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- In large dimension, it is often ore sensible to use hybrid algorithm, to update a subset of the parameters... the eigendecomposition can help there to find the directions which are worthwhile to update.

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- In large dimension, it is often ore sensible to use hybrid algorithm, to update a subset of the parameters... the eigendecomposition can help there to find the directions which are worthwhile to update.
- In presence of non-linear correlation π, estimating a single covariance matrix is not enough. In this case, non-linear ACP methods (e.g. locally linear) are better suited...

Adaptative Metropolis-Hastings Algorithm

Metropolis-Hastings with independent proposals

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- Similar to the A/R algorithm, efficient if the proposal q is close to π...

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 - 1. easy to sample
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- Objective: on-line adaptation of the parameter by minimizing the Kullback divergence

$$\mathrm{KL}(\pi \| q_{\theta}) = \int \log(\frac{\pi(x)}{q_{\theta}(x)}) \pi(x) dx \; .$$

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Method: On-line EM algorithm (see ICASSP 2006)

Adaptative Metropolis-Hastings Algorithm



Figure: Banana shaped target distribution

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Adaptative Metropolis-Hastings Algorithm



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Adaptative Metropolis-Hastings Algorithm



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Adaptative Metropolis-Hastings Algorithm

Results (Andrieu & Moulines, 2006)

 Law of Large Numbers (under assumptions that do not imply the cvge of θ_k)

$$n^{-1} \sum_{k=1}^{n} \left[f(X_k) - \pi(f) \right] \xrightarrow{\mathsf{a.s}}_{\bar{\mathbf{P}}_{\star}} 0 .$$

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• Central Limit Theorem (if $\lim_k \theta_k$ exists)

$$n^{-1/2} \sum_{k=1}^{n} [f(X_k) - \pi(f)] \xrightarrow{\mathcal{D}}_{\bar{\mathbf{P}}_{\star}} Z ,$$

with Z characteristic function $\bar{\mathbf{E}}_{\star} \left[\exp(-\frac{1}{2}\sigma^2(\theta_{\infty},f)t^2) \right]$ and $\sigma^2(\theta_{\infty},f)$ variance of the MCMC under θ_{∞}^*

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An Application



Figure: Global monitoring of gaseous matters (ozone layer) and aerosol concentrations by occultation of stars

An Application



Figure: Principle of the measurement of the transmittance spectrum

An Application



Figure: Spectrum of the star for considered wavelengths

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An Application



Figure: Atmospheric transmittance at different tangential altitude (height). Sirius

An Application

Model

Principle: T(λ, z) = exp (-∑_g α_g(λ)N_g(z)) (Beer & Lambert)
 1. T(λ, z) transmittance at λ and tangential altitude z

An Application

Model

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 - 1. $T(\lambda,z)$ transmittance at λ and tangential altitude z
 - 2. $N_g(z) (\text{mol/cm}^2)$ integrated quantity of gaseous matter (O₃, H₂O, NO₂ ...) at tangential height z. Related to the concentration $z \mapsto \rho_g(z)$ by

$$N_g(z) = \int_{\ell(z)}
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An Application

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3. $\alpha_g(\lambda)$ absorption coefficient of gaseous species g at frequency λ .

An Application

▶ Altitude discretization (approx. 1 km) and $\rho_g(z)$ assumed constant for altitude diff. less than the step-size:

$$N_g(z_i) = \sum_{j=1}^J \ell_{i,j} R_{g,j} \quad R_{g,j} = \rho_g(z_j)$$

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▶ Prior model for the concentration: Gaussian Linear State Space Model, i.e. R_{g,j} = [01]X_{g,j}

$$X_{g,j} = \begin{pmatrix} 1 & 0 \\ 1 & 1 \end{pmatrix} X_{g,j-1} + \begin{pmatrix} \sigma \\ 0 \end{pmatrix} \mathcal{N}(0,1)$$

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An Application

Measurements

Measurements: noisy estimates of the transmittance at frequencies λ₁,..., λ_I et d'altitudes z₁,..., z_J

$$y(\lambda_i, z_j) = T(\lambda_i, z_j) + \varepsilon(\lambda_i, z_j) ,$$

where $\varepsilon(\lambda_i, z_j)$ measurement noise (independent, Gaussian, known variance)...

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► Objective: Infer the posterior distribution of the gaseous component concentration {R_{g,j}, j = 1,...,J, g = 1,...,G}... Well-posed Non-Linear Inverse Problem!

An Application

Main Characteristics

► Huge number of measurements: I ≈ 1500 fréquencies, J ≈ 100 height: 150000 measurement for a single occultation experiments (and up to 10 occultation experiment / day)...

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An Application

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- ► Huge number of measurements: I ≈ 1500 fréquencies, J ≈ 100 height: 150000 measurement for a single occultation experiments (and up to 10 occultation experiment / day)...
- Huge number of variables $J \times G \approx 500$.
- Variability of the experimental set-up star emission spectrum, atmospheric turbulence, line of sight ...

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Adaptation is vital !

-An Application



An Application



Figure: Joint and marginal distributions of two gaseous components at 24 and 26 $\rm km$

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An Application

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Sensible criterion

An Application

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Sensible criterion ↔ understand the chain dynamic (simulation bottleneck)

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- There are many possible ways to adapt a simulation strategy. Most often, it is more difficult to find appropriate adaptation criteria rather than to design the on-line procedure itself.
- Sensible criterion ↔ understand the chain dynamic (simulation bottleneck) ↔ asymptotic analysis (dimension, fluid limit, etc.)

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3. Coupling MCMC (serial) and particle (parallel) methods.

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- 5. most needed ToolBox (AdapBUGS !)

Some theoretical results

Ingredients

• $(P_{\theta}, \theta \in \Theta)$ a family of transition kernels with target distribution π .

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- ▶ $H: \Theta \times X \to \Theta$ an estimating function: for all $\theta \in \Theta$,

$$h(\theta) \stackrel{\text{def}}{=} \iint_{\mathsf{X}} H(x,\theta) \pi(dx) \; .$$

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Algorithm:

$$X_{k+1} \sim P_{\theta_k}(X_k, \cdot)$$

$$\theta_{k+1} = \theta_k + \gamma_{k+1} H(\theta_k, X_{k+1})$$

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Some theoretical results

Problems and Questions.

 $\{(X_k, \theta_k)\}$ is a non-homogeneous Markov Chain but... $\{X_k\}$ is not a Markov Chain ! Q: Is it still ergodic ?

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1. Limit Theorems for Additive Functionals

$$n^{-\gamma} \sum_{k=1}^{n} \left(\psi_{\theta_k}(X_k) - \int_{\mathsf{X}} \psi_{\theta_k}(x) \pi(dx) \right)$$

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2. Rate of Convergence (?)

$$\|\mathbf{E}_{(x,\theta)}[f(X_k)] - \pi(f)\|_{\mathrm{TV}} \le C \|f\|_{\infty} r(k)$$

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Ergodicity is not Automatically preserved...



Figure: Metropolis algorithm on \mathbb{R} . Target $\pi = \mathcal{N}(0, 1)$, Proposal $q = \mathcal{N}(0, \theta^2)$. Adaptation: $\theta^2 = \theta_+^2$ if $X_k \ge 0$ et $\theta^2 = \theta_-^2$ if $X_k < 0$.

Assumptions: Geometric Ergodicity

There exists a function $V: X \to [1, \infty]$ and a set C such that for all $K \subset \Theta$ compact,

► Foster-Lyapunov: $\sup_{\theta \in \mathsf{K}} P_{\theta} V \leq \lambda_{\mathsf{K}} V + b_{sfK} \mathbb{1}_{\mathsf{C}}$

[†]Coupling + Sharpening the Lindvall inequality () +

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(Douc & M.,2003)[†] There exist constants $\rho_{\rm K} < 1$ and $C_{\rm K} < \infty$, depending explicitly on $\lambda_{\rm K}$, $b_{\rm K}$ and $\delta_{\rm K}$, such that

 $\sup_{\theta \in \mathcal{K}} \|P_{\theta}^{n}f - \pi(f)\|_{V} \leq C_{\mathsf{K}} \|f\|_{V} \rho_{\mathsf{K}}^{n} \quad \|f\|_{V} = \sup |f(x)|/V(x)$

[†]Coupling + Sharpening the Lindvall inequality (-)

Some theoretical results

Assumptions: Smoothness

For all $\mathsf{K} \subset \Theta$ compact, $(\theta, \theta') \in \mathsf{K} \times \mathsf{K}$ $\blacktriangleright ||P_{\theta}f - P_{\theta'}f||_{V} \leq C_{\mathsf{K}} ||f||_{V} |\theta - \theta'|$

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(Andrieu & M.,2005) Existence of a solution \hat{f}_{θ} to the Poisson equation $f - \pi(f) = \hat{f}_{\theta} - P_{\theta}\hat{f}_{\theta}$; \hat{f}_{θ} is Lipshitz $\|\hat{f}_{\theta} - \hat{f}_{\theta'}\|_{V} \leq C_{\mathsf{K}}|\theta - \theta'|$ for all $(\theta, \theta') \in \mathsf{K} \times \mathsf{K}$. Satisfied by most Metropolis-Hastings algorithms, (but not always easy).

Error decomposition

Existence of solutions to Poisson Eqs. \Rightarrow

$$f(X_k) - \pi(f) = \hat{f}_{\theta_k}(X_k) - P_{\theta_k}\hat{f}_{\theta_k}(X_k),$$

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First term: martingale. Second term: Lipshitz (cvgce de θ_k not necessary for LLN). Third term: disappear in the summations...

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Assumptions: Stability of the adaptation procedure and convergence

Lyapunov function $w: \Theta \rightarrow [0,\infty]$ such that

1. Level-sets $\mathcal{W}_M \stackrel{\text{def}}{=} \{\theta \in \Theta, w(\theta) \leq M\} \subset \Theta$ are compact,
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(Andrieu & M. et Priouret, 2005) convergence of stochastic approximation $d(\theta_k, \mathcal{L}) \rightarrow 0$ p.s. under verifiable assumptions

Behind Markov Chain Monte-Carlo.

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Some theoretical results

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Stochastic Approximation: an introduction

Let Θ be the domain of allowable values for a vector of parameters θ.

► Two fundamental problems of interest: Problem 1. Find the value(s) of a vector θ ∈ Θ that minimize a scalar-valued loss function w(θ)

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Stochastic Approximation: an introduction

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- Two fundamental problems of interest:
 - Problem 1. Find the value(s) of a vector $\theta \in \Theta$ that minimize a scalar-valued loss function $w(\theta)$
 - Problem 2. Find the value(s) of $\theta \in \Theta$ that solve the equation $h(\theta) = 0$ for some vector-valued function h. Frequently (but not necessarily) $h(\theta) = \nabla w(\theta)$

Stochastic root-finding problem

Focus is on finding θ (i.e., θ*) such that h(θ*) = 0 where h(θ) is typically a nonlinear function of θ assuming that only noisy measurements of h(θ) are available

 $\theta_{k+1} = \theta_k + \gamma_{k+1} Y_{k+1}$ $Y_{k+1} = h(\theta_k) + \text{"noise"}$

- Above problem arises frequently in practice
 - Optimization with noisy measurements (h(θ) represents gradient of loss function)

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- Machine learning
- ... and adaptive MCMC.

Behind Markov Chain Monte-Carlo.

-Some theoretical results

Existence solutions to the Poisson Equation

▶ For any compact subset $\mathcal{K} \subset \Theta$ and for any $r \in [0, 1]$ there exist constants C and $\rho < 1$ such that for all $\psi \in \mathcal{L}_{V^r}$ and all $\theta \in \mathcal{K}$

$$||P_{\theta}^{k}\psi - \pi(\psi)||_{V^{r}} \leq C\rho^{k} ||\psi||_{V^{r}}.$$

▶ Therefore, for all $\theta, x \in \Theta \times X$ and $\psi \in \mathcal{L}_{V^r}$,

$$\sum_{k=0}^{\infty} |P_{\theta}^{k}\psi(x) - \pi(\psi)| < \infty$$

and

$$u \stackrel{\text{def}}{=} \sum_{k=0}^{\infty} (P_{\theta}^{k} \psi - \pi(\psi))$$

is a solution of Poisson's equation: $u - P_{\theta}u = \psi - \pi(\psi)$.

Regularity of the Solutions to the Poisson Equation

for any function f and k,

$$P_{\theta}^{k}f - P_{\theta'}^{k}f = \sum_{j=1}^{k-1} P_{\theta}^{j} \left(P_{\theta} - P_{\theta'}\right) P_{\theta'}^{k-j-1}f$$
$$= \sum_{j=1}^{k-1} P_{\theta}^{j} \left(P_{\theta} - P_{\theta'}\right) \left(P_{\theta'}^{k-j-1}f - \pi(f)\right)$$
$$= \sum_{j=1}^{k-1} \left(P_{\theta}^{j} - \pi\right) \left(P_{\theta} - P_{\theta'}\right) \left(P_{\theta'}^{k-j-1}f - \pi(f)\right)$$

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Behind Markov Chain Monte-Carlo.

Some theoretical results

Regularity of the Solutions to the Poisson Equation

The geometric ergodicity and the regularity of the transition kernel implies that

$$\begin{aligned} |P_{\theta}^{k}f - P_{\theta'}^{k}f| &\leq C_{\theta,\theta'} \sum_{j=1}^{k-1} \rho_{\theta}^{j} \rho_{\theta'}^{k-j} |\theta - \theta'| \\ &\leq C_{\theta,\theta'} \rho^{k} |\theta - \theta'| \quad \rho = \rho_{\theta} \wedge \rho_{\theta'} . \end{aligned}$$

► Denote by \hat{f}_{θ} the solution of the Poisson equation $\hat{f}_{\theta} - P_{\theta}\hat{f}_{\theta} = f - \pi(f)$. Then

$$\hat{f}_{\theta} - \hat{f}_{\theta'} = \sum_{k=1}^{\infty} (P_{\theta}^k f - P_{\theta'}^k f)$$

and the regularity follows from the preceding bound.