Bandits and Exploration

(and a few MDPs)

Tor Lattimore



Contents

- What and why of bandit problems
- A little statistics
- How to solve bandit problems
- Scaling up to RL

Bandits

- Reinforcement learning $S_1, A_1, R_1, S_2, A_2, R_2, \ldots$
- **Bandits** $A_1, R_1, A_2, R_2, \ldots$

Learning is important

Balancing exploration/exploitation important

No planning





Bandits

Finite action set $\mathcal{A} = \{1, 2, \dots, k\}$

For each $a \in \mathcal{A}$ there is an **unknown** distribution P_a

Learner chooses $A_t \in \mathcal{A}$ and observes **reward** $R_t \sim P_{A_t}$

Learner wants to maximise $\sum_{t=1}^{n} R_t$



The learning objective

Let μ_a be the mean of P_a and $\mu^* = \max_{a \in \mathcal{A}} \mu_a$

The optimal action is $a^* = \operatorname{argmax}_a \mu_a$

Our task is to minimise the **regret**

$$\mathfrak{R}_n = n\mu^* - \mathbb{E}\left[\sum_{t=1}^n R_t\right]$$

The price paid by the learner for not knowing μ

A little step into statistics Given independent and identically distributed X, X_1, X_2, \ldots, X_n with mean μ and variance σ^2

The **empirical mean** is
$$\hat{\mu} = rac{1}{n}\sum_{t=1}^n X_t$$

A little step into statistics Given independent and identically distributed X, X_1, X_2, \ldots, X_n with mean μ and variance σ^2

The **empirical mean** is
$$\hat{\mu} = rac{1}{n}\sum_{t=1}^n X_t$$

What does the distribution of μ look like?

We know $\mathbb{E}[\hat{\mu}] = \mu$ and $\operatorname{Var}[\hat{\mu}] = \sigma^2/n$

Chebyshev's inequality:

$$\mathbb{P}\left(\left|\hat{\mu}-\mu\right|\geq\varepsilon\right)\leq\frac{\sigma^2}{n\varepsilon^2}$$



Subgaussian random variables The moment generating function of *X* is

$$M_X(\lambda) = \mathbb{E}[\exp(\lambda X)]$$

A random variable is σ -subgaussian if

$$M_X(\lambda) \le \exp(\sigma^2 \lambda^2/2)$$
 for all $\lambda \in \mathbb{R}$

 $\begin{array}{ll} \mbox{Gaussian} & X \sim \mathcal{N}(\mu,\sigma^2) & X-\mu \mbox{ is } \sigma\mbox{-subgaussian} \\ \mbox{Bernoulli} & X \sim \mathcal{B}(\mu) & X-\mu \mbox{ is } \frac{1}{2}\mbox{-subgaussian} \end{array}$

 $\mathbb{P}\left(\hat{\mu} - \mu \ge \varepsilon\right)$

 $\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \varepsilon\right) = \inf_{\lambda > 0} \mathbb{P}\left(\exp\left(\lambda(\hat{\mu} - \mu)\right) \ge \exp(\lambda\varepsilon)\right)$$

 $\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \varepsilon\right) = \inf_{\lambda > 0} \mathbb{P}\left(\exp\left(\lambda(\hat{\mu} - \mu)\right) \ge \exp(\lambda\varepsilon)\right)$$
$$\leq \inf_{\lambda > 0} \exp(-\lambda\varepsilon)\mathbb{E}\left[\exp\left(\lambda(\hat{\mu} - \mu)\right)\right]$$

 $\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$

 $\mathbb{P}\left(|Z| \ge c\right) \le \mathbb{E}[|Z|]/c$

$$\mathbb{P}(\hat{\mu} - \mu \ge \varepsilon) = \inf_{\lambda > 0} \mathbb{P}(\exp(\lambda(\hat{\mu} - \mu)) \ge \exp(\lambda\varepsilon))$$

$$\leq \inf_{\lambda > 0} \exp(-\lambda\varepsilon) \mathbb{E}\left[\exp(\lambda(\hat{\mu} - \mu))\right]$$

$$= \inf_{\lambda > 0} \exp(-\lambda\varepsilon) \prod_{t=1}^{n} \mathbb{E}\left[\exp\left(\frac{\lambda(X_t - \mu)}{n}\right)\right]$$

$$\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$$

$$\lambda(\hat{\mu} - \mu) = \sum_{t=1}^{n} \frac{\lambda(X_t - \mu)}{n}$$

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \varepsilon\right) = \inf_{\lambda > 0} \mathbb{P}\left(\exp\left(\lambda(\hat{\mu} - \mu)\right) \ge \exp(\lambda\varepsilon)\right)$$

$$\leq \inf_{\lambda > 0} \exp(-\lambda\varepsilon) \mathbb{E}\left[\exp\left(\lambda(\hat{\mu} - \mu)\right)\right]$$

$$= \inf_{\lambda > 0} \exp(-\lambda\varepsilon) \prod_{t=1}^{n} \mathbb{E}\left[\exp\left(\frac{\lambda(X_t - \mu)}{n}\right)\right]$$

$$\leq \inf_{\lambda > 0} \exp(-\lambda\varepsilon) \prod_{t=1}^{n} \exp\left(\frac{\sigma^2\lambda^2}{2n^2}\right)$$

$$\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$$

$$\lambda(\hat{\mu} - \mu) = \sum_{t=1}^{n} \frac{\lambda(X_t - \mu)}{n}$$

$$\begin{split} \mathbb{P}\left(\hat{\mu}-\mu\geq\varepsilon\right) &= \inf_{\lambda>0}\mathbb{P}\left(\exp\left(\lambda(\hat{\mu}-\mu)\right)\geq\exp(\lambda\varepsilon)\right)\\ &\leq \inf_{\lambda>0}\exp(-\lambda\varepsilon)\mathbb{E}\left[\exp\left(\lambda(\hat{\mu}-\mu)\right)\right]\\ &= \inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\mathbb{E}\left[\exp\left(\frac{\lambda(X_t-\mu)}{n}\right)\right]\\ &\leq \inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\exp\left(\frac{\sigma^2\lambda^2}{2n^2}\right)\\ &= \inf_{\lambda>0}\exp\left(\frac{\sigma^2\lambda^2}{2n}-\lambda\varepsilon\right) \end{split}$$

$$\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$$

$$\lambda(\hat{\mu} - \mu) = \sum_{t=1}^{n} \frac{\lambda(X_t - \mu)}{n}$$

$$\begin{split} \mathbb{P}\left(\hat{\mu}-\mu\geq\varepsilon\right) &= \inf_{\lambda>0}\mathbb{P}\left(\exp\left(\lambda(\hat{\mu}-\mu)\right)\geq\exp(\lambda\varepsilon)\right)\\ &\leq \inf_{\lambda>0}\exp(-\lambda\varepsilon)\mathbb{E}\left[\exp\left(\lambda(\hat{\mu}-\mu)\right)\right]\\ &= \inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\mathbb{E}\left[\exp\left(\frac{\lambda(X_t-\mu)}{n}\right)\right]\\ &\leq \inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\exp\left(\frac{\sigma^2\lambda^2}{2n^2}\right)\\ &= \inf_{\lambda>0}\exp\left(\frac{\sigma^2\lambda^2}{2n}-\lambda\varepsilon\right) \end{split}$$

$$\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$$

$$0 = \frac{d}{d\lambda} \left(\frac{\sigma^2 \lambda^2}{2n} - \lambda \varepsilon \right) = \lambda \sigma^2 / n - \varepsilon$$

$$\begin{split} \mathbb{P}\left(\hat{\mu}-\mu\geq\varepsilon\right) &=\inf_{\lambda>0}\mathbb{P}\left(\exp\left(\lambda(\hat{\mu}-\mu)\right)\geq\exp(\lambda\varepsilon)\right)\\ &\leq\inf_{\lambda>0}\exp(-\lambda\varepsilon)\mathbb{E}\left[\exp\left(\lambda(\hat{\mu}-\mu)\right)\right]\\ &=\inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\mathbb{E}\left[\exp\left(\frac{\lambda(X_t-\mu)}{n}\right)\right]\\ &\leq\inf_{\lambda>0}\exp(-\lambda\varepsilon)\prod_{t=1}^{n}\exp\left(\frac{\sigma^2\lambda^2}{2n^2}\right)\\ &=\inf_{\lambda>0}\exp\left(\frac{\sigma^2\lambda^2}{2n}-\lambda\varepsilon\right) =\exp\left(-\frac{n\varepsilon^2}{2\sigma^2}\right) \end{split}$$

$$0 = \frac{d}{d\lambda} \left(\frac{\sigma^2 \lambda^2}{2n} - \lambda \varepsilon \right) = \lambda \sigma^2 / n - \varepsilon$$

 $\exp(\lambda(X-\mu)) \le \exp(\lambda^2 \sigma^2/2)$

Last slide we proved that

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \varepsilon\right) \le \exp\left(-\frac{n\varepsilon^2}{2\sigma^2}\right)$$

Equating the right-hand side with δ and rearranging things a little,

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \sqrt{\frac{2\sigma^2 \log(1/\delta)}{n}}\right) \le \delta$$

for any $\delta \in (0,1).$ Chebyshev's only gives

$$\mathbb{P}\left(\hat{\mu} - \mu \ge \sqrt{\frac{\sigma^2}{n\delta}}\right) \le \delta$$

Concentration of measure summary

- Understanding the **distribution** of the **empirical mean** is important
- Without assumptions **Chebyshev's** is about the best you can do
- Subgaussian assumption leads to much stronger results
- Method is called Chernoff's method
- There are whole books on this topic



Assumptions

We assume $X - \mu_a$ is 1-subgaussian when $X \sim P_a$ for all actions

Subgaussian bandits

Optimism principle

"You should act as if you are in the **nicest plausible** world possible"



Optimism principle

"You should act as if you are in the **nicest plausible** world possible"



Guarantees either (a) optimality or (b) exploration

"Nicest" In bandits, we want the mean to be large

"Plausible" The mean cannot be *much* larger than the empirical mean

"Nicest" In bandits, we want the mean to be large

"Plausible" The mean cannot be *much* larger than the empirical mean

Upper Confidence Bound Algorithm
Choose each arm once and then
$$A_t = \operatorname{argmax}_a \hat{\mu}_a(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}}$$

 $\hat{\mu}_a(t) =$ empirical mean of arm a after round t $T_a(t) =$ number of plays of arm a after round t $\delta =$ confidence level

Regret analysis

- **Step 1** Decompose the regret over the arms
- **Step 2** On a "good" event prove that suboptimal arms are not played too often
- **Step 3** Show the "good" event occurs with high probability

$$\mathfrak{R}_n = n\mu^* - \mathbb{E}\left[\sum_{t=1}^n R_t\right]$$

$$\Delta_a = \mu^* - \mu_a$$
$$T_a(t) = \sum_{s=1}^t \mathbb{1}(A_s = a)$$

$$\mathfrak{R}_n = n\mu^* - \mathbb{E}\left[\sum_{t=1}^n R_t\right]$$
$$= \mathbb{E}\left[\sum_{t=1}^n (\mu^* - R_t)\right]$$

$$\Delta_a = \mu^* - \mu_a$$
$$T_a(t) = \sum_{s=1}^t \mathbb{1}(A_s = a)$$

$$\mathfrak{R}_n = n\mu^* - \mathbb{E}\left[\sum_{t=1}^n R_t\right]$$
$$= \mathbb{E}\left[\sum_{t=1}^n (\mu^* - R_t)\right]$$
$$= \mathbb{E}\left[\sum_{t=1}^n \Delta_{A_t}\right]$$

$$\Delta_a = \mu^* - \mu_a$$
$$T_a(t) = \sum_{s=1}^t \mathbb{1}(A_s = a)$$



$$\Delta_a = \mu^* - \mu_a$$
$$T_a(t) = \sum_{s=1}^t \mathbb{1}(A_s = a)$$

$$\begin{aligned} \mathfrak{R}_n &= n\mu^* - \mathbb{E}\left[\sum_{t=1}^n R_t\right] \\ &= \mathbb{E}\left[\sum_{t=1}^n (\mu^* - R_t)\right] \\ &= \mathbb{E}\left[\sum_{t=1}^n \Delta_{A_t}\right] \\ &= \mathbb{E}\left[\sum_{t=1}^n \sum_{a \in \mathcal{A}} \mathbb{1}(A_t = a)\Delta_a\right] \\ &= \sum_{a \in \mathcal{A}} \Delta_a \mathbb{E}[T_a(n)] \end{aligned}$$

$$\Delta_a = \mu^* - \mu_a$$
$$T_a(t) = \sum_{s=1}^t \mathbb{1}(A_s = a)$$

Assume for all t that $\hat{\mu}$

$$\hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \ge \mu^*$$
$$\mu_a + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \ge \hat{\mu}_a(t-1)$$

Assume for all *t* that

$$\hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \ge \mu^*$$
$$\mu_a + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \ge \hat{\mu}_a(t-1)$$

Now suppose that $A_t = a$ in round t

$$\mu_a + 2\sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \ge \hat{\mu}_a(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}}$$

Assume for all *t* that

$$\hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \ge \mu^*$$
$$\mu_a + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \ge \hat{\mu}_a(t-1)$$

Now suppose that $A_t = a$ in round t

$$\begin{aligned} \mu_a + 2\sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} &\geq \hat{\mu}_a(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \\ &\geq \hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \geq \mu_{a^*} \end{aligned}$$

Assume for all *t* that

. .

$$\hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \ge \mu^*$$
$$\mu_a + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \ge \hat{\mu}_a(t-1)$$

Now suppose that $A_t = a$ in round t

$$\begin{aligned} \mu_a + 2\sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} &\geq \hat{\mu}_a(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_a(t-1)}} \\ &\geq \hat{\mu}_{a^*}(t-1) + \sqrt{\frac{2\log(1/\delta)}{T_{a^*}(t-1)}} \geq \mu_{a^*} \end{aligned}$$

Hence
$$T_a(t-1) \le \frac{8\log(1/\delta)}{\Delta_a^2} \implies T_a(n) \le 1 + \frac{8\log(1/\delta)}{\Delta_a^2}$$

Let $\hat{\mu}_{a,s}$ be the empirical mean of arm a after s plays

The concentration theorem shows that

$$\mathbb{P}\left(\hat{\mu}_{a,s} \ge \mu_a + \sqrt{\frac{2\log(1/\delta)}{s}}\right) \le \delta$$

Combining with a union bound,

$$\mathbb{P}\left(\text{exists } s \le n : \hat{\mu}_{a,s} \ge \mu_a + \sqrt{\frac{2\log(1/\delta)}{s}}\right) \le n\delta$$

 $\mathbb{P}\left(\cup_{i}B_{i}\right)\leq\sum_{i}\mathbb{P}\left(B_{i}\right)$

Putting it together

$$\begin{aligned} \mathfrak{R}_n &= \sum_{a \in \mathcal{A}} \Delta_a \mathbb{E}[T_a(n)] \\ &\leq \sum_{a \in \mathcal{A}: \Delta_a > 0} \Delta_a \left(2\delta n^2 + 1 + \frac{8\log(1/\delta)}{\Delta_a^2} \right) \\ &\leq \sum_{a \in \mathcal{A}: \Delta_a > 0} 3\Delta_a + \frac{16\log(n)}{\Delta_a} \end{aligned}$$

Sanity checking our results

We have proven the regret of UCB is at most


$$\mathfrak{R}_n = \sum_{a \in \mathcal{A}} \Delta_a \mathbb{E}[T_a(n)]$$

$$\Re_n = \sum_{a \in \mathcal{A}} \Delta_a \mathbb{E}[T_a(n)]$$
$$= \sum_{a \in \mathcal{A}: \Delta_a \leq \Delta} \Delta_a \mathbb{E}[T_a(n)] + \sum_{a \in \mathcal{A} \Delta_a > \Delta} \Delta_a \mathbb{E}[T_a(n)]$$

$$\begin{aligned} \mathfrak{R}_{n} &= \sum_{a \in \mathcal{A}} \Delta_{a} \mathbb{E}[T_{a}(n)] \\ &= \sum_{a \in \mathcal{A}: \Delta_{a} \leq \Delta} \Delta_{a} \mathbb{E}[T_{a}(n)] + \sum_{a \in \mathcal{A} \Delta_{a} > \Delta} \Delta_{a} \mathbb{E}[T_{a}(n)] \\ &\leq n\Delta + \sum_{a \in \mathcal{A}: \Delta_{a} > \Delta} 3\Delta_{a} + \frac{16 \log(n)}{\Delta_{a}} \end{aligned}$$

$$\begin{aligned} \mathfrak{R}_{n} &= \sum_{a \in \mathcal{A}} \Delta_{a} \mathbb{E}[T_{a}(n)] \\ &= \sum_{a \in \mathcal{A}: \Delta_{a} \leq \Delta} \Delta_{a} \mathbb{E}[T_{a}(n)] + \sum_{a \in \mathcal{A} \Delta_{a} > \Delta} \Delta_{a} \mathbb{E}[T_{a}(n)] \\ &\leq n\Delta + \sum_{a \in \mathcal{A}: \Delta_{a} > \Delta} 3\Delta_{a} + \frac{16 \log(n)}{\Delta_{a}} \\ &\leq n\Delta + \frac{16K \log(n)}{\Delta} + 3\sum_{a \in \mathcal{A}} \Delta_{a} \end{aligned}$$

$$\begin{aligned} \Re_n &= \sum_{a \in \mathcal{A}} \Delta_a \mathbb{E}[T_a(n)] \\ &= \sum_{a \in \mathcal{A}: \Delta_a \leq \Delta} \Delta_a \mathbb{E}[T_a(n)] + \sum_{a \in \mathcal{A} \Delta_a > \Delta} \Delta_a \mathbb{E}[T_a(n)] \\ &\leq n\Delta + \sum_{a \in \mathcal{A}: \Delta_a > \Delta} 3\Delta_a + \frac{16\log(n)}{\Delta_a} \\ &\leq n\Delta + \frac{16K\log(n)}{\Delta} + 3\sum_{a \in \mathcal{A}} \Delta_a \\ &\leq 8\sqrt{nk\log(n)} + 3\sum_{a \in \mathcal{A}} \Delta_a \leq 8\sqrt{nk\log(n)} + 3k \end{aligned}$$

There is a lot more..

- Improving constants
- Different noise models
- Linear bandits: $\mathcal{A} \subset \mathbb{R}^d$ and $\mu_a = \langle \mu, a \rangle$
- Other kinds of structure: $\mathcal{A} \subset \mathbb{R}^d$ and $\mu_a = f(a)$ with f 'smooth'
- Changing action sets
- Delayed rewards
- Non-stationary bandits
- Best arm identification
- Adversarial model

Lots of fun still to be had, but this is an RL workshop

Exploration in reinforcement learning ("We want states")

Episodic MDPs

An **episodic MDP** is a tuple (S, A, P, H, r, μ)

- + ${\mathcal S}$ is a finite set of ${\it states}$
- \mathcal{A} is a finite set of **actions**
- $\cdot P$ is the transition kernel
- \cdot *H* is the **episode length**
- $r: \mathcal{S} \times \mathcal{A} \rightarrow [0, 1]$ is the **reward function**
- + μ is the distribution of the initial state

Episodic MDPs

An **episodic MDP** is a tuple (S, A, P, H, r, μ)

- $\cdot \ \mathcal{S}$ is a finite set of **states**
- \mathcal{A} is a finite set of **actions**
- P is the transition kernel
- \cdot *H* is the **episode length**
- $r: \mathcal{S} \times \mathcal{A} \rightarrow [0, 1]$ is the **reward function**
- + μ is the distribution of the initial state

Assumption Only *P* is unknown

 $S = \{1, 2, 3\}$ and H = 4



Policies and values

A **policy** π is a function from histories to actions The **value** of a policy π is

$$v^{\pi} = \mathbb{E}\left[\sum_{h=1}^{H} r(S_h, A_h)\right]$$

Dynamic programming Think of $P(s, a) = (P(s, a, 1), \dots, P(s, a, |\mathcal{S}|))$

The optimal value function is defined inductively

 $v_0(s) = 0$ $q_h(s, a) = r(s, a) + \langle P(s, a), v_{h-1} \rangle$ $v_h(s) = \max_{a \in \mathcal{A}} q_h(s, a)$ $\pi_h(s) = \operatorname{argmax}_{a \in \mathcal{A}} q_h(s, a)$

$$\mathcal{P} = \{ x \in [0,1]^{|\mathcal{S}|} : \|x\|_1 = 1 \}$$

Learning and regret

- In each episode the learner chooses a policy π^t
- Observes a trajectory $S_1^t, A_1^t, S_2^t, A_2^t, \dots, S_H^t, A_H^t$
- Regret over n episodes is

$$\Re_n = \sum_{t=1}^n \Re^{(t)} = \mathbb{E}\left[\sum_{t=1}^n \langle \mu, v_H^* - v_H^{\pi^t} \rangle\right]$$

Optimism for RL

- Same idea!
- Estimate the things you don't know (transitions)
- Build confidence intervals around the unknowns
- Act as if the world is as nice as plausible

Estimation and confidence intervals

The empirical transitions are given by

 $T_{s,a}(t) =$ # plays action a in state s $\hat{P}_t(s, a, s') =$ # prop. transitions to s' from s taking a

Estimation and confidence intervals

The empirical transitions are given by

$$T_{s,a}(t) = \#$$
 plays action a in state s
 $\hat{P}_t(s, a, s') = \#$ prop. transitions to s' from s taking a

The confidence set is ℓ_1 -ball about vector $\hat{P}_t(s, a)$

$$\mathcal{C}_t(s,a) = \left\{ p \in \mathcal{P} : \left\| p - \hat{P}_t(s,a) \right\|_1 \le \sqrt{\frac{2|\mathcal{S}|\log(2/\delta)}{T_{s,a}(t)}} \right\}$$

$$\mathcal{P} = \{ x \in [0,1]^{|\mathcal{S}|} : \|x\|_1 = 1 \}$$

Optimistic dynamic programming

At the start of phase t,

 $\tilde{v}_0(s) = 0$ $\tilde{q}_h(s,a) = r(s,a) + \max_{p \in \mathcal{C}_{t-1}(s,a)} \langle p, \tilde{v}_{h-1} \rangle$ $\tilde{v}_h(s) = \max_{a \in \mathcal{A}} \tilde{q}_h(s, a)$ $\pi_h^t(s) = \operatorname{argmax}_{a \in \mathcal{A}} \tilde{q}_h(s, a)$ $P_h(s) = \operatorname{argmax}_{p \in \mathcal{C}_{t-1}(s, \pi_h(s))} \langle p, \tilde{v}_{h-1} \rangle$

UCB for reinforcement learning

Three steps in each episode

Step 1 Compute empirical estimate of transitions and confidence intervals

Step 2 Use optimistic dynamic programming to find a policy

Step 3 Implement policy for entire episode

Algorithm is called **Upper Confidence Bounds for Reinforcement** Learning (UCRL)

Analysing UCRL

Use optimism

With high probability $P(s, a) \in C_t(s, a)$ for all t and s, a

Analysing UCRL

Use optimism

With high probability $P(s, a) \in C_t(s, a)$ for all t and s, a

Assuming this holds, then

$$\begin{split} \langle \mu, v_H - v_H^{\pi^t} \rangle &= \langle \mu, v_H \rangle - \langle \mu, v_H^{\pi^t} \rangle \\ &\leq \langle \mu, \tilde{v}_H^{\pi^t} \rangle - \langle \mu, v_H^{\pi^t} \rangle \\ &= \langle \mu, \tilde{v}_H^{\pi^t} - v_H^{\pi^t} \rangle \end{split}$$

Useful because it's **much** easier to compare values under the same policy

Value differences

Decompose the value difference:

$$\langle \mu, \tilde{v}_{H}^{\pi^{t}} - v_{H}^{\pi^{t}} \rangle = \mathbb{E} \left[\sum_{h=1}^{H} \langle \tilde{P}_{H-h+1}^{t}(S_{h}^{t}, A_{h}^{t}) - P(S_{h}^{t}, A_{h}^{t}), \tilde{v}_{H-h}^{\pi^{t}} \rangle \right]$$

We might look at the proof later

$$\mathfrak{R}^{(t)} \lesssim \mathbb{E}\left[\sum_{h=1}^{H} \langle \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h), \tilde{v}_{H-h}^{\pi} \rangle\right]$$

Hölder's inequality: $\langle x, y \rangle \leq \|x\|_1 \|y\|_{\infty}$

$$\mathfrak{R}^{(t)} \lesssim \mathbb{E}\left[\sum_{h=1}^{H} \langle \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h), \tilde{v}_{H-h}^{\pi} \rangle\right]$$
$$\leq \mathbb{E}\left[\sum_{h=1}^{H} \left\| \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h) \right\|_1 \left\| \tilde{v}_{H-h}^{\pi} \right\|_{\infty}\right]$$

Hölder's inequality: $\langle x, y \rangle \leq ||x||_1 ||y||_{\infty}$

$$\begin{aligned} \mathfrak{R}^{(t)} &\lesssim \mathbb{E}\left[\sum_{h=1}^{H} \langle \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h), \tilde{v}_{H-h}^{\pi} \rangle \right] \\ &\leq \mathbb{E}\left[\sum_{h=1}^{H} \left\| \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h) \right\|_1 \| \tilde{v}_{H-h}^{\pi} \|_{\infty} \right] \\ &\lesssim H \mathbb{E}\left[\sum_{h=1}^{H} \sqrt{\frac{|\mathcal{S}| \log(1/\delta)}{T_{S_h, A_h}(t-1)}} \right] \end{aligned}$$

Hölder's inequality: $\langle x, y \rangle \leq \|x\|_1 \|y\|_{\infty}$

$$\begin{aligned} \mathfrak{R}^{(t)} &\lesssim \mathbb{E} \left[\sum_{h=1}^{H} \langle \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h), \tilde{v}_{H-h}^{\pi} \rangle \right] \\ &\leq \mathbb{E} \left[\sum_{h=1}^{H} \left\| \tilde{P}_{H-h+1}(S_h, A_h) - P(S_h, A_h) \right\|_1 \| \tilde{v}_{H-h}^{\pi} \|_{\infty} \right] \\ &\lesssim H \mathbb{E} \left[\sum_{h=1}^{H} \sqrt{\frac{|\mathcal{S}| \log(1/\delta)}{T_{S_h, A_h}(t-1)}} \right] \\ &\lesssim H \mathbb{E} \left[\sum_{s, a} T_{s, a}(t-1, t) \sqrt{\frac{|\mathcal{S}| \log(1/\delta)}{T_{s, a}(t-1)}} \right] \end{aligned}$$

Hölder's inequality: $\langle x,y\rangle \leq \|x\|_1\,\|y\|_\infty$

$$\sum_{t=1}^{n} \mathfrak{R}^{(t)} \leq H\mathbb{E} \left[\sum_{s,a} \sum_{t=1}^{n} T_{s,a}(t-1,t) \sqrt{\frac{|\mathcal{S}| \log(1/\delta)}{T_{s,a}(t-1)}} \right]$$
$$\lesssim H\mathbb{E} \left[\sum_{s,a} \sqrt{|\mathcal{S}| T_{s,a}(n) \log(1/\delta)} \right]$$
$$\leq H\mathbb{E} \left[\sqrt{|\mathcal{S}|^2 |\mathcal{A}| \sum_{s,a} T_{s,a}(n) \log(1/\delta)} \right]$$
$$= H|\mathcal{S}| \sqrt{|\mathcal{A}| Hn \log(1/\delta)}$$

$$\int \frac{f'(x)}{\sqrt{f(x)}} dx = 2\sqrt{f(x)}$$

At last...

With 'high probability' the regret of UCRL is

$$\mathfrak{R}_n = O\left(|\mathcal{S}|H\sqrt{n|\mathcal{A}|\log(1/\delta)}\right)$$

Lower bound Any algorithm has regret at least

$$\mathfrak{R}_n = \Omega\left(H\sqrt{n|\mathcal{A}||\mathcal{S}|\log(1/\delta)}\right)$$

Takeaways

- A little concentration of measure
- Optimism as a principle for algorithm design
- Optimism for bandits (UCB) and MDPs (UCRL)

Let us reflect for a moment

Let us reflect for a moment How big is $H\sqrt{n|\mathcal{A}||\mathcal{S}|\log(1/\delta)}$?



Let us reflect for a moment How big is $H\sqrt{n|\mathcal{A}||\mathcal{S}|\log(1/\delta)}$?



$$|\mathcal{S}| = 2^{20}$$

Let us reflect for a moment How big is $H\sqrt{n|\mathcal{A}||\mathcal{S}|\log(1/\delta)}$?





Big challenges

- Exploring in large unstructured MDPs is hopeless
- Combining exploration with function approximation
- Bringing in bias
- Optimism is not universal
- All known exploration principles are either (a) known to be suboptimal or (b) hopelessly intractible
- Model free exploration

Great time to be in RL (theory and practice!)

"Bandit Algorithms" book

Joint work with Csaba Szepesvári

Free online at http://banditalgs.com





Reading

- UCB. Tze Leung Lai. Adaptive Treatment Allocation and the Multi-Armed Bandit Problem, 1987
- UCRL. Auer et al. Near-optimal Regret Bounds for Reinforcement Learning, 2010

Useful keywords Posterior sampling, information directed sampling, Bellman rank, randomized value functions. Preface with 'deep' for more buzz

Categorical concentration

Let X, X_1, X_2, \ldots, X_n be independent and identically distributed with $X_t \in [k]$

Let
$$p_i = \mathbb{P}(X = i)$$
 and $\hat{p}_i = \frac{1}{n} \sum_{t=1}^n \mathbb{1}(X_t = i)$

You can have fun proving that

$$\mathbb{P}\left(\|p-\hat{p}\|_1 \geq \sqrt{\frac{2k\log(2/\delta)}{n}}\right) \leq \delta$$