

Nonnegative garrote in additive models using P-splines

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JOINT WORK WITH ANESTIS ANTONIADIS AND IRÈNE GIJBELS

Outline

1. Nonnegative garrote in additive model
2. P-spline estimator
3. Smoothing spline estimator
4. Backfitting
5. Consistency
6. COSSO-type P-spline estimator
7. Simulations and Boston housing data

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Nonnegative garrote

- **Multiple linear model:** response Y , d explanatory variables X_1, \dots, X_d

$$Y_i = \sum_{j=1}^d \beta_j X_{ij} + \varepsilon_i$$

- **Nonnegative garrote (Breiman, 1995):**

$$\begin{cases} (\hat{c}_1, \dots, \hat{c}_d) = \operatorname{argmin}_{\vec{c}} \sum_{i=1}^n \left(Y_i - \sum_{j=1}^d c_j \hat{\beta}_j^{\text{LS}} X_{ij} \right)^2 \\ \text{s.t. } c_j \geq 0 \ (j = 1, \dots, d), \quad \sum_{j=1}^d c_j \leq s \end{cases}$$

\iff

$$\begin{cases} (\hat{c}_1, \dots, \hat{c}_d) = \operatorname{argmin}_{\vec{c}} \left\{ \sum_{i=1}^n \left(Y_i - \sum_{j=1}^d c_j \hat{\beta}_j^{\text{LS}} X_{ij} \right)^2 + \lambda \sum_{j=1}^d c_j \right\} \\ \text{s.t. } c_j \geq 0 \end{cases}$$

$$\implies \hat{\beta}_j^{\text{NNG}} = \hat{c}_j \hat{\beta}_j^{\text{LS}}$$

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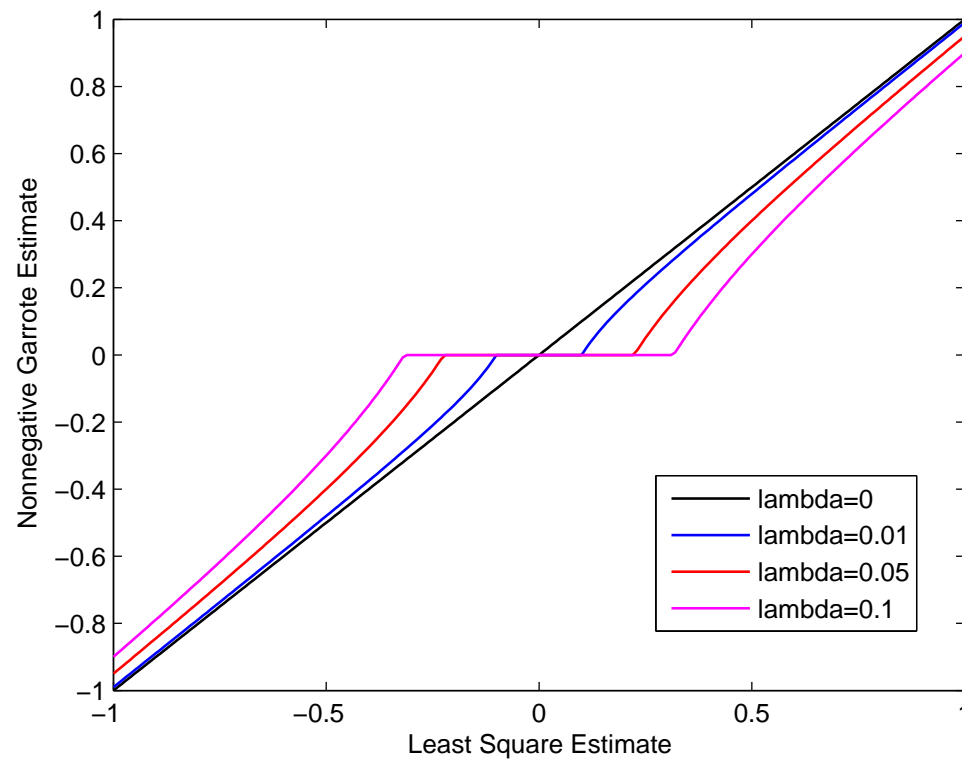
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Orthogonal design: $X'X = I_{n \times n}$

$$\hat{\beta}_j^{\text{NNG}} = \left(1 - \frac{\lambda}{2(\hat{\beta}_j^{\text{LS}})^2}\right)_+ \hat{\beta}_j^{\text{LS}}$$



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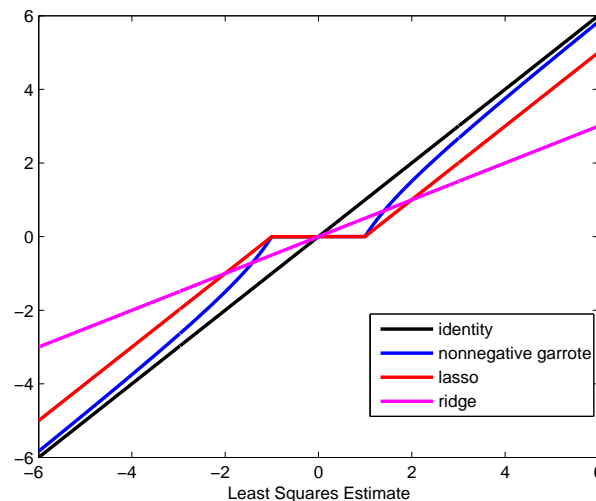


Lasso:

$$\begin{cases} (\hat{\beta}_1^{\text{Lasso}}, \dots, \hat{\beta}_d^{\text{Lasso}}) = \operatorname{argmin}_{\vec{\beta}} \sum_{i=1}^n \left(Y_i - \sum_{j=1}^d \beta_j X_{ij} \right)^2 \\ \text{s.t. } \sum_{j=1}^d |\beta_j| \leq s \end{cases}$$

Ridge:

$$\begin{cases} (\hat{\beta}_1^{\text{Ridge}}, \dots, \hat{\beta}_d^{\text{Ridge}}) = \operatorname{argmin}_{\vec{\beta}} \sum_{i=1}^n \left(Y_i - \sum_{j=1}^d \beta_j X_{ij} \right)^2 \\ \text{s.t. } \sum_{j=1}^d \beta_j^2 \leq s \end{cases}$$



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Functional nonnegative garrote

- **Additive model:** response Y , d explanatory variables X_1, \dots, X_d

$$Y_i = \sum_{j=1}^d f_j(X_{ij}) + \varepsilon_i$$

- **Functional nonnegative garrote** (Cantoni et al., 2006 and Yuan, 2007):

$$\left\{ \begin{array}{l} (\hat{c}_1, \dots, \hat{c}_d) = \operatorname{argmin}_{\vec{c}} \sum_{i=1}^n \left(Y_i - \sum_{j=1}^d c_j \hat{f}_j^{\text{init}}(X_{ij}) \right)^2 \\ \text{s.t. } c_j \geq 0 \ (j = 1, \dots, d), \quad \sum_{j=1}^d c_j \leq s \end{array} \right.$$

$$\implies \hat{f}_j^{\text{NNG}} = \hat{c}_j \hat{f}_j^{\text{init}}$$

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P-spline estimator

- Regression of n data points (X_i, Y_i) , with

$$Y_i = \mu(X_i) + \varepsilon_i$$

ε_i i.i.d. $N(0, \sigma^2)$

- Representation of $\mu(\cdot)$ in a q -th degree B-spline basis:

$$\mu(x) = \sum_{j=1}^m \alpha_j B_j(x; q)$$

- Penalized least squares estimator $\hat{\alpha}$: minimizer of

$$\sum_{i=1}^n \left(Y_i - \sum_{j=1}^m \alpha_j B_j(X_i; q) \right)^2 + \lambda \sum_{j=k+1}^m (\Delta^k \alpha_j)^2 = \|Y - B\vec{\alpha}\|_2^2 + \lambda \|D_k \vec{\alpha}\|_2^2$$

For example: $\Delta^2 \alpha_j = \alpha_j - 2\alpha_{j-1} + \alpha_{j-2}$

$$\implies \text{P-spline estimator } \hat{\mu}(x) = \sum_{j=1}^m \hat{\alpha}_j B_j(x; q)$$



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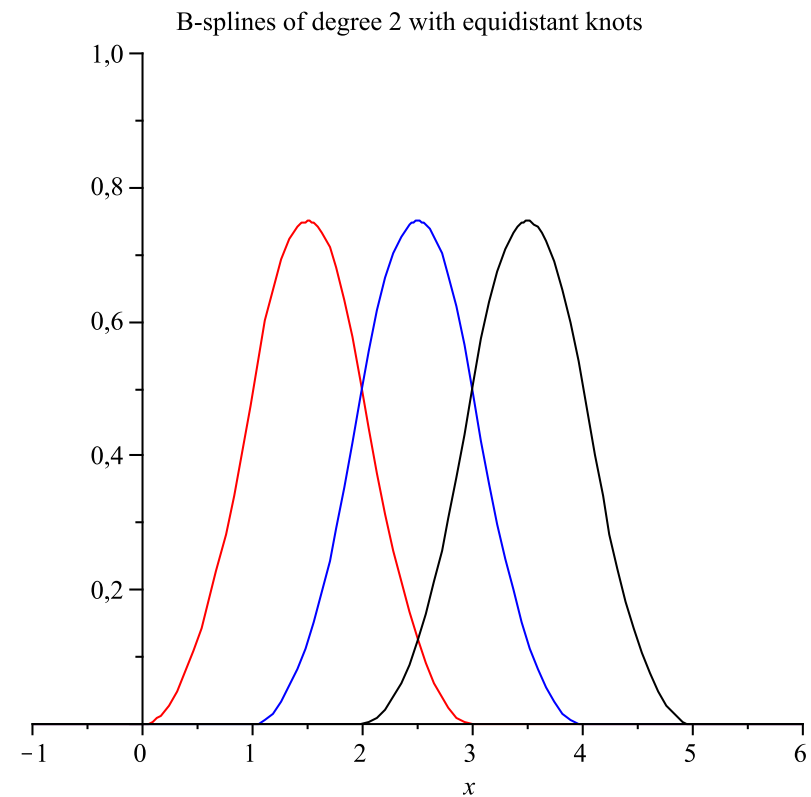
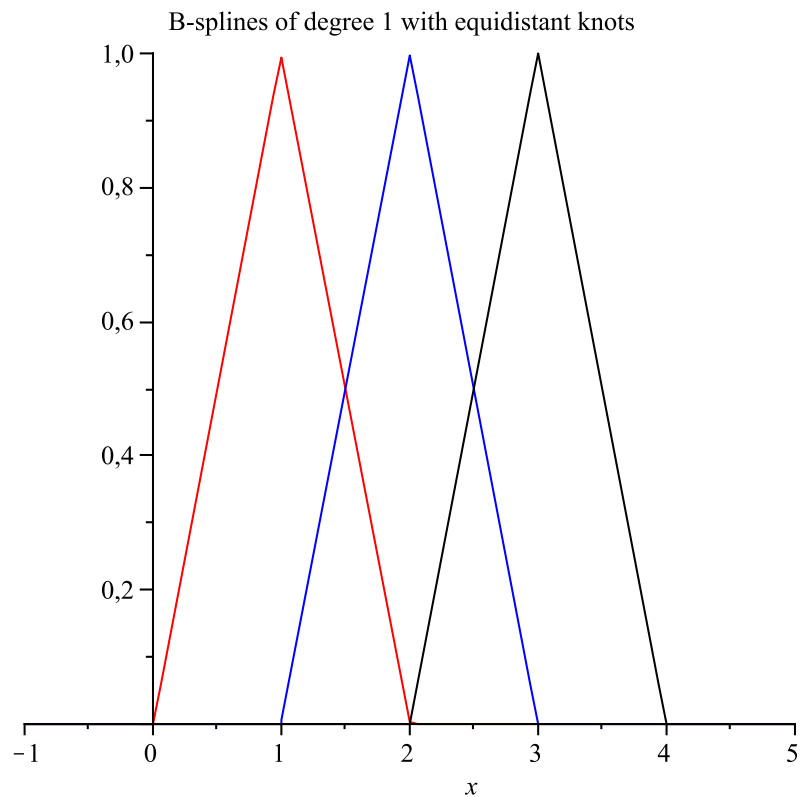
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Smoothing spline estimator

- Regression of n data points (X_i, Y_i) , with

$$Y_i = \mu(X_i) + \varepsilon_i$$

ε_i i.i.d. $N(0, \sigma^2)$

- Smoothing spline estimator $\hat{\mu}(x)$: minimizer of

$$\sum_{i=1}^n \left(Y_i - \mu(X_i) \right)^2 + \lambda \int (\mu^{(k)}(x))^2 dx$$

- If $k = 2$, $\hat{\mu}(x)$ is a cubic spline on $[X_{(1)}, \dots, X_{(n)}]$



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Backfitting

$$Y_i = \sum_{j=1}^d f_j(X_{ij}) + \varepsilon_i$$

Algorithm:

1. Set $\hat{f}_j(X_{1j}, \dots, X_{nj}) = 0$ for $j = 1, \dots, d$
2. Repeat steps 3 to 5 until the estimates $\hat{f}_j(X_{1j}, \dots, X_{nj})$ stop changing
3. For $j = 1, \dots, d$ repeat steps 4 and 5
4. Calculate the partial residuals

$$e_j = Y - \sum_{h \neq j} \hat{f}_h(X_{1h}, \dots, X_{nh})$$

5. Set $\hat{f}_j(X_{1j}, \dots, X_{nj})$ equal to the result of smoothing e_j with respect to x_j

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- **Yuan (2007):** Assume that for some $\delta_n \rightarrow 0$

$$\frac{1}{n} \sum_{i=1}^n (f_j(x_{ij}) - \hat{f}_j^{\text{init}}(x_{ij}))^2 = \frac{1}{n} \|f_j - \hat{f}_j^{\text{init}}\|_2^2 = O_P(\delta_n^2).$$

If $\lambda \rightarrow 0$ in a fashion such that $\delta_n = o_P(\lambda)$ then

$$\sup_j \frac{1}{n} \|f_j - \hat{f}_j^{\text{NNG}}\|_2^2 = O_P(\lambda^2) \text{ and } P(\hat{f}_j^{\text{NNG}} = 0) \rightarrow 1 \text{ if } f_j = 0$$

- **Claeskens, Krivobokova and Opsomer (2008):**

$$\frac{1}{n} \sum_{i=1}^n \mathbb{E}((f_j(x_{ij}) - \hat{f}_j(x_{ij}))^2) = \frac{1}{n} \mathbb{E}(\|f_j - \hat{f}_j\|_2^2) = O(n^{-\frac{2q+2}{2q+3}})$$

- **Horowitz, Klemelä and Mammen (2006):**

each additive component can be estimated as well as it could be if the others were known

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COSSO-type P-spline estimator

- **COSSO estimator** (Lin and Zhang, 2006): function $f \in \mathcal{F}$ (with $\mathcal{F} = \{1\} \oplus_{j=1}^d \mathcal{F}_j$ a RKHS) that minimizes

$$\sum_{i=1}^n (Y_i - \sum_{j=1}^d f_j(X_{ij}))^2 + \lambda \sum_{j=1}^d \|P^j f\|_{\mathcal{F}}$$

- **COSSO-type P-spline estimator** $(\hat{\alpha}_1, \dots, \hat{\alpha}_d)$: minimizer of

$$\|Y - \sum_{j=1}^d B_j \vec{\alpha}_j\|_2^2 + \sum_{j=1}^d \lambda_j \|D_k \vec{\alpha}_j\|_2$$

$$\implies \hat{f}(x_1, \dots, x_d) = \sum_{j=1}^d \sum_{h=1}^m \hat{\alpha}_{hj} B_h(x_j; q)$$



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● **COSSO-type P-spline estimator**

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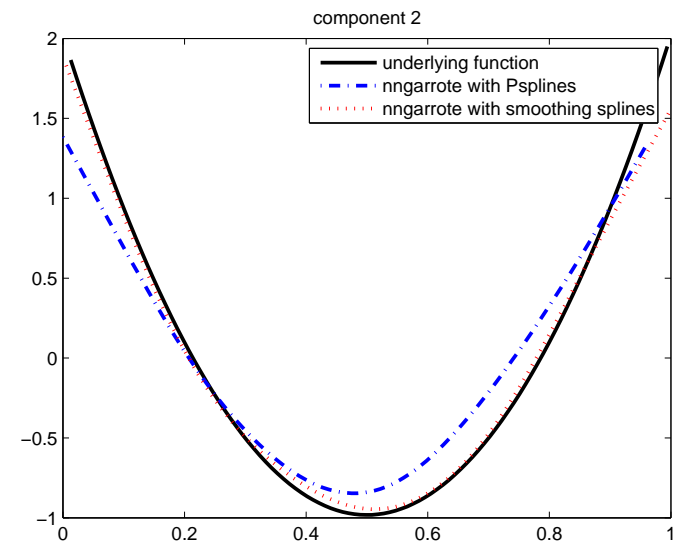
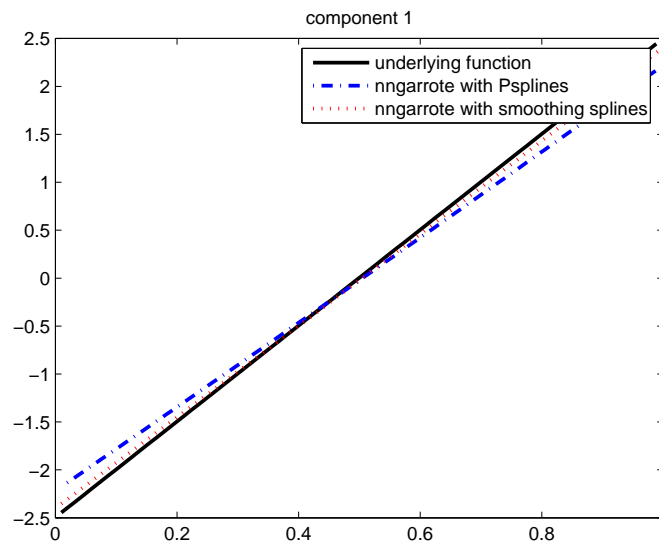
Simulated data

$$Y_i = f_1(X_{i1}) + f_2(X_{i2}) + f_3(X_{i3}) + f_4(X_{i4}) + \varepsilon_i$$

$$f_1(x) = 5x, \quad f_2(x) = 3(2x - 1)^2, \quad f_3(x) = \frac{\sin(2\pi x)}{2 - \sin(2\pi x)}$$

$$f_4(x) = 0.6 \sin(2\pi x) + 1.2 \cos(2\pi x) + 1.8 \sin^2(2\pi x) + 2.4 \cos^3(2\pi x) + 3 \sin^3(2\pi x)$$

$$AISE_j = \frac{1}{n} \sum_{i=1}^n (\hat{f}_j^{\text{NNG}}(X_{ij}) - f_j(X_{ij}))^2$$



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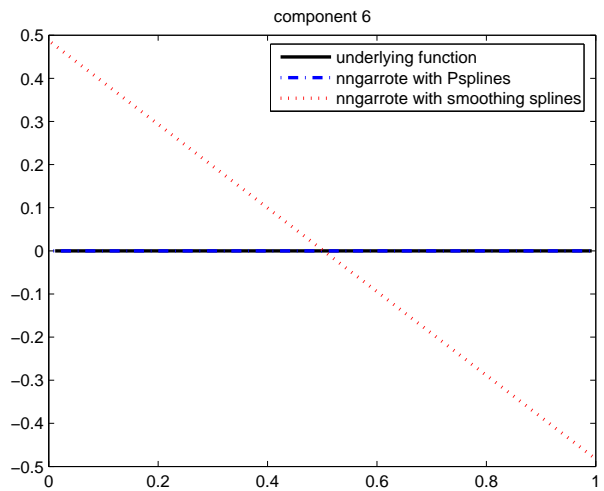
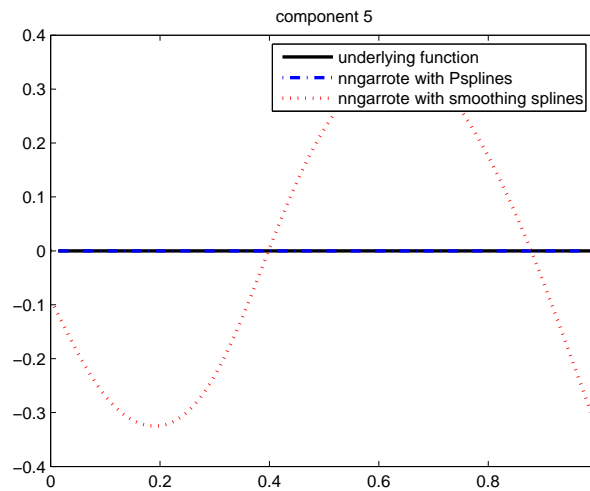
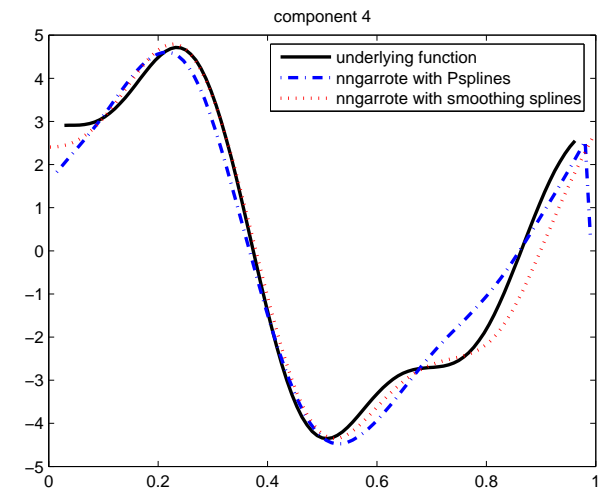
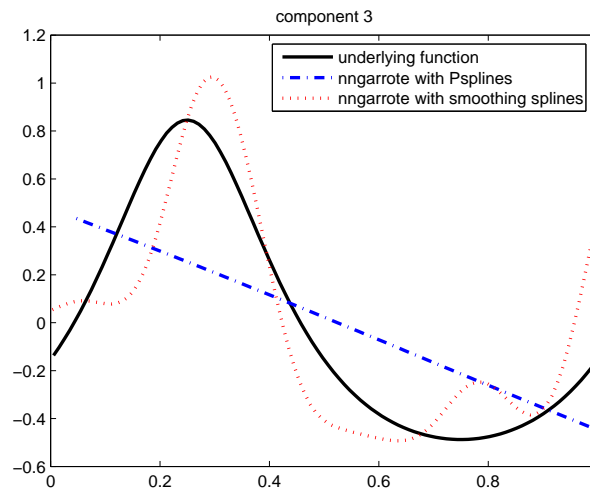
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Boston Housing data



$$\begin{aligned} \log(\text{medv}) = & \alpha + f_1(\text{crim}) + f_2(\text{zn}) + f_3(\text{indus}) + f_4(\text{nox}) + f_5(\text{rm}) \\ & + f_6(\text{age}) + f_7(\text{dis}) + f_8(\text{rad}) + f_9(\text{tax}) + f_{10}(\text{ptratio}) \\ & + f_{11}(\text{b}) + f_{12}(\text{lstat}) + \beta \text{chas} + \varepsilon \end{aligned}$$

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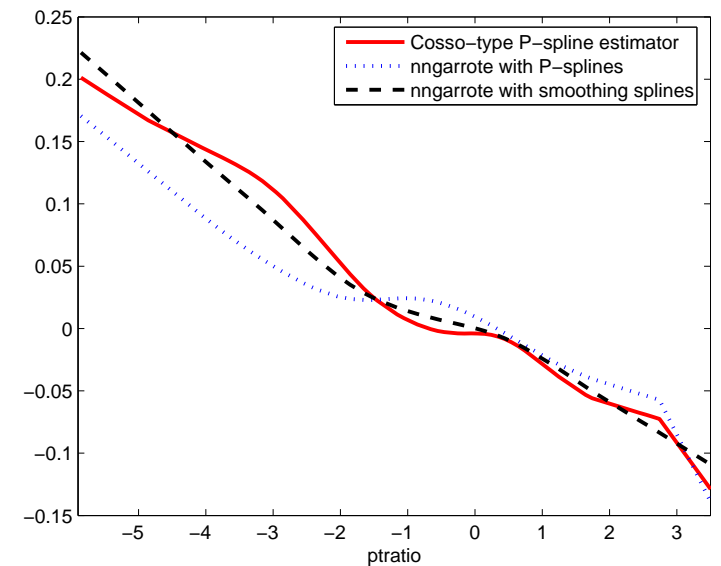
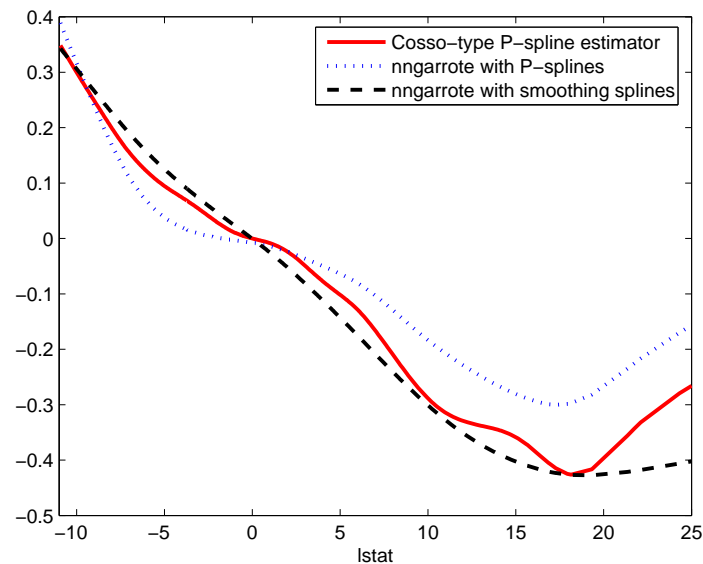
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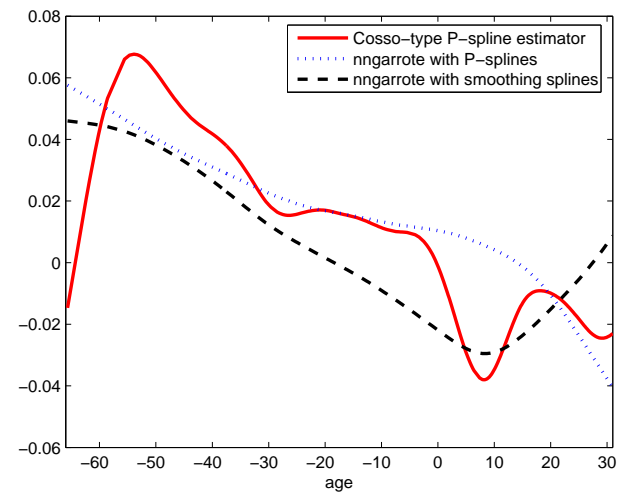
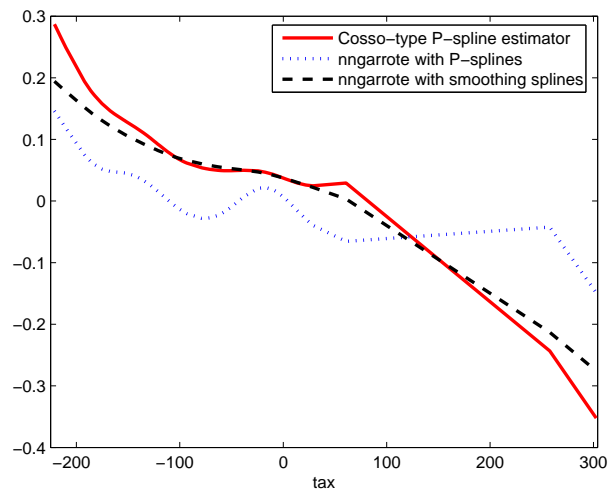
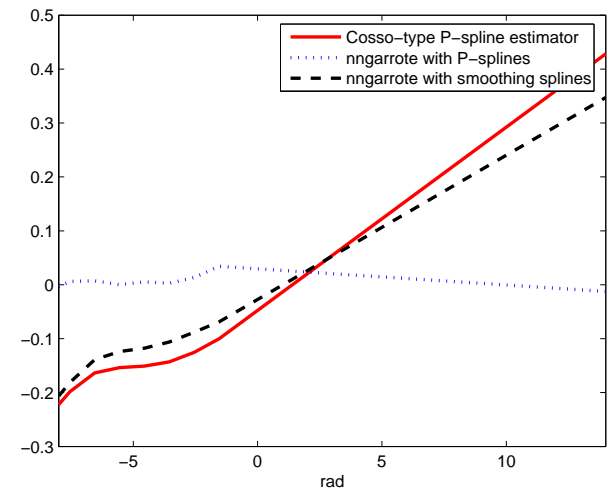
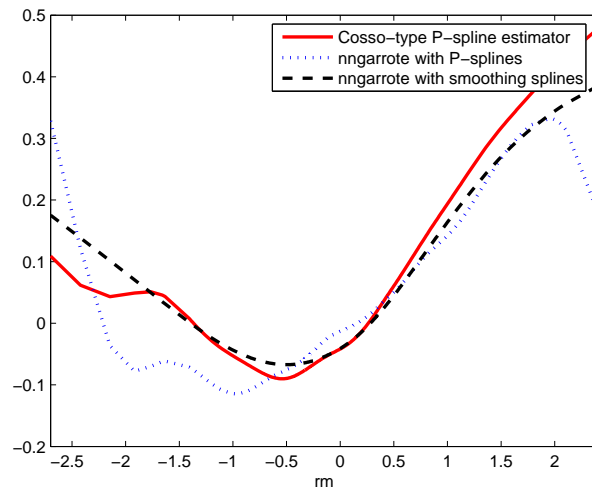
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References

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