

# Moment closure and block updating for parameter inference in stochastic biological models

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## 1 Motivation

- Auto regulatory gene network

## 2 Method

- Moment closure
- Approximate Likelihood
- Updating missing data

## 3 Results

- Data set  $D_1$  -  $G$  and  $I$  unobserved
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- Conclusions and future work

# Motivation

- One of the key problems in systems biology is inferring rate parameters of stochastic kinetic biochemical network models
- If we know:
  - 1 The description of the system
  - 2 The initial conditions
  - 3 The rate parameters

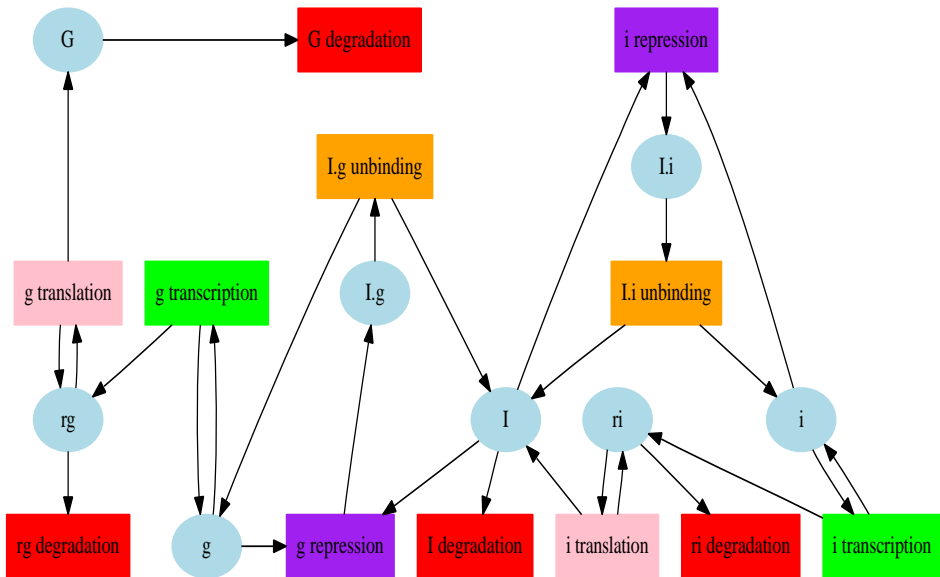
then we can model the system (stochastically or deterministically)

- Test our understanding of the system/modelling assumptions
- How do we infer these rate parameters initially?

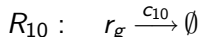
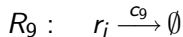
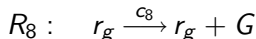
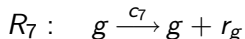
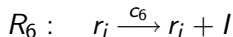
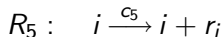
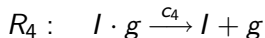
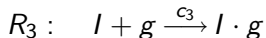
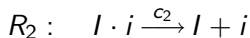
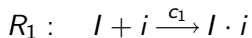
# Auto regulatory gene network

Throughout this talk we will use an auto regulatory gene network as an example

- This network has 6 species  $Z = (r_g, r_i, g, i, G, I)$
- $r_g$  and  $r_i$  are mRNA
- $g$  and  $i$  are genes
- $G$  and  $I$  are proteins
- Where  $I$  regulates the production of itself and  $G$  by binding to genes  $i$  and  $g$



- We can write our model as a list of reactions:



- We assume mass action kinetics

- From the chemical master equation we can find a set of ODE's for the moments (see Gillespie, 2009)
- ODE's for the moments usually depend on higher order moments, e.g. for two species  $X_1, X_2$

$$\dot{\mu}_{1,1} = (\mu_{2,0} - \mu_{1,0}^2)c_1 - (\mu_{1,1} - \mu_{1,0}\mu_{0,1})c_1 - \mu_{2,1}c_1 + \dots$$

where  $\mu_{n,m} = E(X_1^n X_2^m)$

- By assuming an underlying distribution we can write higher order moments in terms of lower order moments e.g.  
 $\mu_3 = 3\mu_2\mu_1 - 2\mu_1^3$
- Giving a closed set of ODE's
- We have assumed a underlying Gaussian distribution throughout this talk, other distributions could be used e.g. Poisson, Log-Normal

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- Let  $\mathbf{x}(t_i)$  be the  $i^{\text{th}}$  discrete time observation of the process
- We propose

$$\mathbf{x}(t_i) | \mathbf{x}(t_{i-1}) \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma}),$$

where  $\boldsymbol{\mu}$  and  $\boldsymbol{\Sigma}$  are calculated from the moment closure approximation of the process

- Appealing to the Markov property we can approximate the likelihood of the rate parameters ( $\Theta$ ) for a given realisation  $\mathbf{x} = \{\mathbf{x}(t_i) : i \in 1, \dots, N\}$ ,

$$L(\Theta | \mathbf{x}) = \prod_{i=1}^N P[\mathbf{x}(t_i) | \mathbf{x}(t_{i-1})]$$

- We use a Metropolis-Hastings sampler to explore the parameter space (random walk with innovations  $w_j \sim N(0, \sigma^2)$ )

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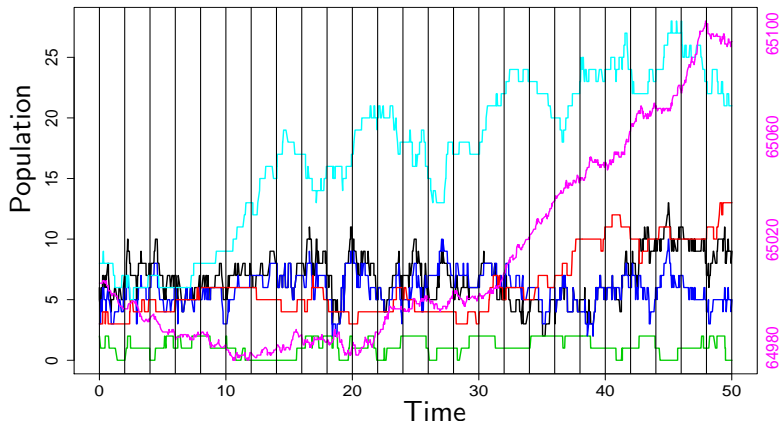
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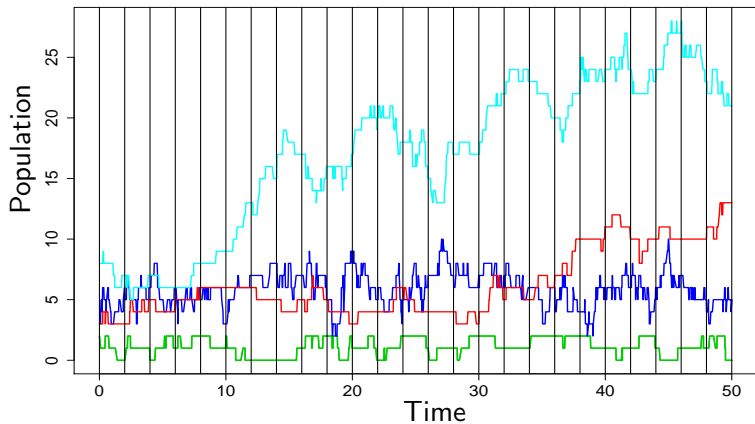
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## Given discrete time observations



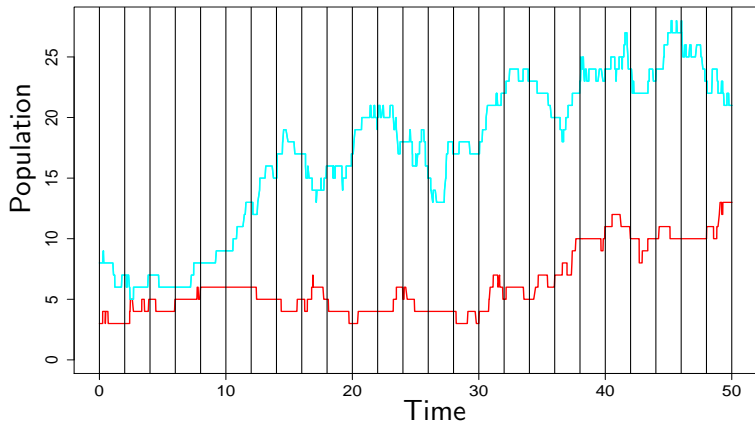
**Figure:** A stochastic realisation from the auto-regulatory gene network. With observations on each species;  $r_g$  (cyan),  $r_i$  (red),  $g$  (blue),  $i$  (green),  $G$  (pink) and  $I$  (black).  $Z(0) = (8, 2, 3, 2, 65000, 6)$

This may be a bit hopeful so we consider  $D_1 = \{r_g, r_i, g, i\}$



**Figure:** A stochastic realisation from the auto-regulatory gene network. With observations on each species;  $r_g$  (cyan),  $r_i$  (red),  $g$  (blue) and  $i$  (green).

This may still be a bit hopeful so we consider  $D_2 = \{r_g, r_i\}$



**Figure:** A stochastic realisation from the auto-regulatory gene network. With observations on each species;  $r_g$  (cyan) and  $r_i$  (red).

# Bridge updating

- How to update the unobserved species?
- We want to be able to update our missing data conditioned on all the data we can
- We do this using a block updating scheme (following Durham & Gallant (2002))

- Suppose we have data  $Z(t) = (X(t), Y(t))^T$ , where  $X(t)$  is known
- Our goal is to sample  $Y(t_{i+1})$  conditioned on  $Z(t_j)$ ,  $Z(t_M)$  and  $X(t_{i+1})$ , where  $t_j < t_{i+1} < t_M$
- Such a sample can be approximated by a skeleton bridge  $Y(t_{i+1})$  for  $i = j, j + 1, \dots, M - 2$
- Constructing such a bridge is non trivial so a Metropolis Hastings step is used

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- We can construct a proposal distribution for  $Y^{i+1}$

$$q(Y^{i+1}|X^{i+1}, Z^i, Z^M, \theta) \sim N \left\{ \mu^*, \frac{M-i-1}{M-i} \Sigma^* \right\},$$

where,

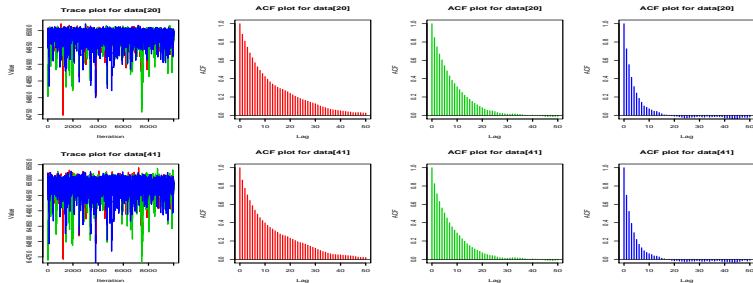
$$\begin{aligned} \mu^* &= \mu_y + \Sigma_{yx}(\Sigma_{xx})^{-1}(X^{i+1} - \mu_x) \\ \Sigma^* &= [\Sigma_{yy} - \Sigma_{yx}(\Sigma_{xx})^{-1}\Sigma_{xy}] \end{aligned}$$

and,

$$\mu_x = X^i + \frac{X^M - X^i}{M-i}, \quad \mu_y = Y^i + \frac{Y^M - Y^i}{M-i}$$

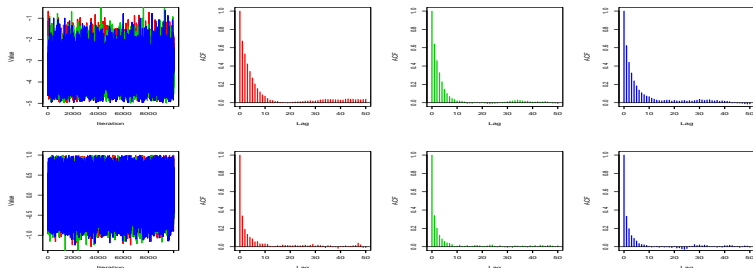
- We can sample  $q(\cdot|\cdot)$  for  $i = j, \dots, M-2$  to construct a skeleton bridge

- We will now apply our block updating method to the two data sets
  - 1  $D_1$ : We have 50 observations on  $X(t) = (r_g, r_i, g, i)$  and impute  $Y(t) = (G, I)$
  - 2  $D_2$ : We have 50 observations on  $X(t) = (r_g, r_i)$  and impute  $Y(t) = (g, i, G, I)$
- In each data set we have limited the number of genes ( $i$ ) to 2 and the steady state value for  $G \approx 70000$ .
- We would like to know:
  - 1 Which block size is best for updating the missing data
  - 2 How much we can find out about our rate parameters and unobserved species



	$M = 1$	$M = 4$	$M = 8$
G[20]	344	571	1417
G[41]	382	623	1547
Mean (all)	465	747	1656

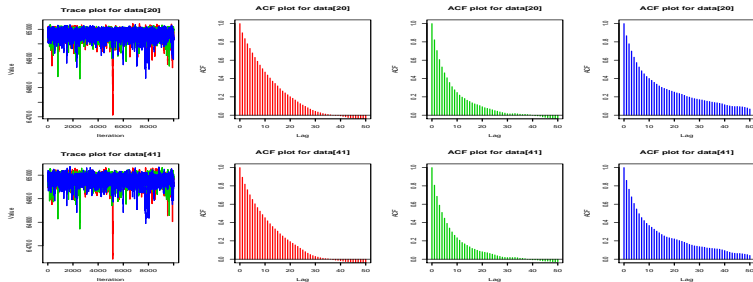
Table: Effective sample sizes for different blocks ( $G$ )



	$M = 1$	$M = 4$	$M = 8$	True	Mean	sd
$c_1$	1305	1936	1830	0.08	0.052	0.043
$c_4$	3125	3618	3624	0.9	0.96	0.50

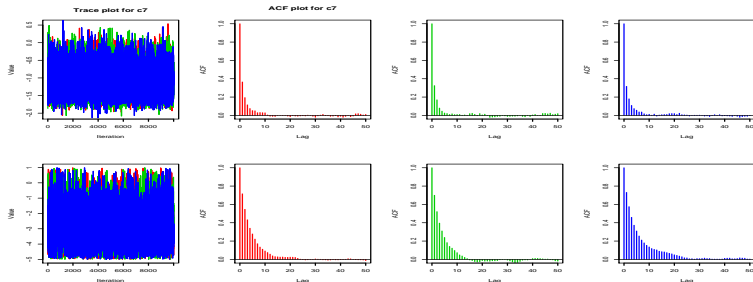
Table: Effective sample sizes of the parameters  $c_1$  and  $c_4$ , for different blocks.

# Data set $D_2 - g, i, G$ and $I$ unobserved



	$M = 1$	$M = 4$	$M = 8$
G[20]	411	731	402
G[41]	413	759	445
Mean (all)	460	912	495

Table: Effective sample sizes for different blocks updating  $G$ .



	$M = 1$	$M = 4$	$M = 8$	True	Mean	sd
$c_7$	3458	4410	3765	0.35	0.37	0.15
$c_{11}$	1266	1431	1146	0.05	0.15	0.24

Table: Effective sample sizes of  $c_7$  and  $c_{11}$  for different blocks.

## Conclusions and future work

- The most efficient choice of block length is model specific, a block length of 4-8 gave the best results in testing
- Conditioning on the observed data leads to more efficient updating of the unobserved data
- Develop a model for *Bacillus subtilis* sporulation and apply our method

## References

- Durham, G. B. & Gallant, R. A. (2002), Numerical techniques for maximum likelihood estimation of continuous time diffusion processes. *Journal of Business and Economic Statistics* **20**, 279-316.
- Gillespie, C. S. (2009), Moment closure approximations for mass-action models. *IET Systems Biology* **3**, 52-58.
- Wilkinson, D. J. (2006). Stochastic Modelling for Systems Biology. *Chapman & Hall/CRC*